

Large-Sample Evidence on Firms' Year-over-Year MD&A Modifications

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ABSTRACT

The Securities and Exchange Commission (SEC) has expressed concern about the informativeness of firms' Management Discussion and Analysis (MD&A) disclosures. A firm's MD&A is potentially uninformative if it does not change appreciably from the previous year after significant economic changes at the firm. We introduce a measure for narrative disclosure—the degree to which the MD&A differs from the previous disclosure—and provide three findings on the usefulness of MD&A disclosure. First, firms with larger economic changes modify the MD&A more than those with smaller economic changes. Second, the magnitude of stock price responses to 10-K filings is positively associated with the MD&A modification score, but analyst earnings forecast revisions are unassociated with the score, suggesting that investors—but not analysts—use MD&A information. Finally, MD&A modification scores have declined in the past decade even as MD&A disclosures have become longer; the price reaction to MD&A modification scores has also weakened, suggesting a decline in MD&A usefulness.

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Data availability: <http://www.accountingresearch.org/data>

1. Introduction

This paper examines the extent to which firms modify the Management Discussion and Analysis (MD&A) from year to year, the usefulness of these modifications to market participants, and the trend in modifications and usefulness over time. As a key narrative disclosure required by the Securities and Exchange Commission (SEC) for annual and quarterly financial reporting, the MD&A must cover certain topics, but managers have flexibility in choosing the breadth and depth of what is discussed. The MD&A is intended as a way for managers to provide their perspectives of the firm to investors, such as why earnings have changed, what liquidity needs the firm faces, what capital resources have been or are planned to be used, and what material market risks the firm is exposed to. Such information would help investors assess the firm's past performance and current financial condition as well as predict future cash flows. An MD&A is unlikely to serve this purpose if the document does not change appreciably from the previous year, especially after significant economic changes.

The modification aspect of MD&A disclosure has also been a concern to regulators. The SEC warned about the lack of modifications in narrative SEC filings in *A Plain English Handbook: How to Create Clear SEC Disclosure Documents*, published in 1998 (SEC [1998], p. 12). For the MD&A in particular, in three places in its 2003 MD&A guidance the SEC urged firms to "evaluate issues presented in previous periods and consider reducing or omitting discussion of those that may no longer be material or helpful or revise discussions where a revision would make the continuing relevance of an issue more apparent."¹ Recently, the SEC told managers to "cut down on the needless information it [Corporate America] discloses to investors" and said it was "looking at its rules to determine whether companies are being asked to provide the 'right' information."² This announcement mentioned the MD&A as a key interest of the SEC.

We develop an MD&A modification score based on an algorithm commonly used by Internet search engines to determine similarities between documents. We use this measure to examine three questions. First, do firms with larger economic changes modify the MD&A more than firms with smaller economic changes? An affirmative finding would suggest that firms meet a baseline level of disclosure.³ Second, do investors and financial analysts use the information revealed by MD&A modifications? The

¹See Part III. A, B.1, and B.2 of "Release No. 33-8350: Commission guidance regarding management's discussion and analysis of financial condition and results of operations" by the SEC on December 19, 2003.

²See "SEC to firms: Cut the 'mind numbing' disclosures," by Rachele Younglai, Reuters, November 4, 2009.

³Although the SEC does not prescribe a minimum (baseline) compliance level, we use the phrase "meet a minimum (baseline) disclosure requirement" to mean that as expected from the regulation, firms with larger economic changes provide more extensive disclosure than those with smaller economic changes.

usefulness of MD&A disclosure has been an important issue to investors and regulators. On the one hand, the MD&A was mandated as a unique venue for providing information to investors and thus the modifications are expected to be useful. On the other hand, managers may use their discretion to strategically obfuscate the financial results (Li [2008]). In addition, the MD&A may not be useful for *real-time* decision making because of its relative lack of timeliness. Third, have managers' MD&A modification behaviors and market participants' reactions to the MD&A changed over time? This question is relevant in light of the significant changes in firms' regulatory and information environments in the past decade.

For the first question—do firms with larger economic changes modify the MD&A more than firms with smaller economic changes?—we use a range of variables to capture a firm's changes in operations, liquidity, capital resources, risk exposure, and business components. We find a positive association between the MD&A modification score and these variables, suggesting that firms meet a minimum disclosure requirement. In addition, managers appear to modify the document to a larger degree after changes in liquidity and capital resources (LCR) than after changes in operations, suggesting that the MD&A is perhaps a better information source about LCR than about earnings. Further, by holding the economic changes constant, we find that cross-sectional variation in MD&A modifications is associated with firm size, competition, and legal environment—factors that have been identified in prior research as determinants of other voluntary disclosure decisions. This finding lends credibility to the idea that the modification score captures information in narrative disclosure.

For the second question—do investors and financial analysts use the information revealed by MD&A modifications?—we use a three-day stock return and find that investors respond to firms' 10-K filings more strongly when the MD&A is modified to a larger degree. The overwhelming majority of financial analysts, however, do not revise their earnings forecasts in the 30 days following the filing date. Further, the magnitude of forecast revisions from the analysts who *do* revise is not associated with the MD&A modification score. These results suggest that equity investors use the MD&A information but analysts do not, perhaps because the MD&A contains new information for investors to predict cash flows in *future periods* but provides little new information about the *subsequent period's* earnings.

To further understand the findings related to the first two questions, we hand-code the MD&A documents of a subsample of firms in the top and bottom quintiles of the modification score. We find that high-score firms more often discuss different aspects of business in the current year (that do not appear in the previous year) than do low-score firms. Even for the aspects that a firm discusses in both years, high-score firms change the discussion to a larger extent than do low-score firms. These results further validate our modification score. More importantly, we observe that the above differences are largely driven by discussion about operations and LCR, not by risk factor disclosures and the use of cautionary language. Furthermore,

we find that managers discuss different aspects of LCR (that do not appear in the previous year) more often than they discuss operations. These observations corroborate our findings of (1) a higher association of MD&A modifications with LCR changes than with earnings changes and (2) the usefulness of MD&A modifications to investors but not to analysts.

Finally, to answer the third question—have firms' MD&A modification behaviors and market participants' reactions to the MD&A changed over time?—we examine the trend of MD&A modification scores and market participants' reactions over the past decade. While MD&A disclosures have become longer over time, they have become more like what investors saw in the previous year. The combined trends of increasing MD&A length and decreasing MD&A modification scores suggest that, over time, managers increasingly use boilerplate disclosure (i.e., standard disclosure that uses many words with little firm-specific or fiscal-period-specific content). Moreover, we find that the price responses to MD&A modifications have weakened over time. These findings suggest a decline in MD&A usefulness in recent years despite the SEC 2003 guidance on improving the MD&A.

Our study makes three contributions to the accounting literature. First, we propose a new measure for narrative disclosure. Measuring narrative disclosure is challenging because the disclosure is qualitative, difficult to quantify, and prone to disagreement among evaluators. Although our measure is not perfect, it is one step forward in understanding and quantifying corporate narrative disclosure. It complements existing measures, such as document length, readability (sentence length and the number of syllables per sentence), tone (positive versus negative words), forward-looking word counts, and small-sample hand coding. Our measure is applicable to many other accounting settings where the disclosure is narrative, its timing is routine, but its content is discretionary, such as proxy statements, annual CEO letters to shareholders, earnings announcement press releases, and the prepared portion of quarterly conference calls.

Second, our study provides fresh evidence of managers' MD&A disclosure behaviors. Compared to earnings announcement press releases, the form of MD&A disclosure is mandatory and its timing is late. Compared to the notes accompanying the financial statements, the content of MD&A disclosure is discretionary and not audited. Given this mixed mandatory/discretionary nature and lack of timeliness, the quality and usefulness of the MD&A have been issues since the MD&A was required in 1980. Constrained by the high costs of hand-processing long narrative documents, the MD&A literature has provided limited evidence of the usefulness of MD&A disclosure (Cole and Jones [2005]). We examine one important facet of MD&A disclosure from a large sample—modifications from the previous year—and provide findings helpful for investors and researchers to understand the usefulness of the MD&A. As the SEC seeks more *informative* narrative disclosure from its filers and the International Accounting Standards Board has advised, but not mandated, that companies prepare an MD&A (IASB [2009]), our findings are also timely and relevant to

regulators. Our third contribution is to extend the emerging large-sample text analysis literature.⁴

The paper proceeds as follows. Section 2 provides background information about MD&A regulation and discusses related research. Section 3 explains how we measure the differences between documents. Section 4 describes the data and provides descriptive statistics. Section 5 tests whether firms with larger economic changes modify the MD&A more than firms with smaller economic changes and examines the firm characteristics associated with cross-sectional variation in MD&A modifications. Section 6 tests investors' and analysts' responses to the MD&A. Section 7 reports our findings from a hand-coded subsample of MD&A documents, section 8 examines the trend of MD&A modifications and market participants' reactions, and section 9 concludes.

2. *Background and Related Research*

2.1 BACKGROUND

Item 303 of Regulation S-K mandates that companies provide an MD&A as Item 7 in the 10-K filing. Managers are required to discuss (1) the results of operations and (2) LCR. Additional required topics now include critical accounting policies and estimates, market risk disclosure, and off-balance-sheet arrangements. The SEC has adopted a management approach to allow investors to "see the company through the eyes of management," helping investors understand *why* the operating results and financial condition have changed and assess the implications of these changes for future cash flows (SEC [2003]). While the discretion accorded to managers allows them to tailor disclosure to suit each business, it also gives managers leeway to keep disclosure to a minimum.

The SEC has conducted several targeted reviews of firms' MD&A practices and provided three interpretive releases to guide MD&A disclosure (SEC [1987], [1989], and [2003]). The most recent review was on the Fortune 500 companies' MD&A disclosures filed in 2002 and, as a result, 350 companies received comment letters.⁵ After the review, the SEC issued the most recent interpretive guidance on December 19, 2003 and emphasized

⁴See Li [2008, 2010], Lerman and Livnat [2010], Feldman et al. [2010], Kothari, Li, and Short [2009], Frankel, Mayew, and Sun [2010], Nelson and Pritchard [2007], Matsumoto, Pronk, and Roelofsen [2007], Davis, Piger, and Sedor [2008], Davis and Tama-Sweet [2009], and Muslu et al. [2009].

⁵"We issued a significant number of comments generally seeking greater analysis of the company's financial condition and results of operations. Our comments addressed situations where companies simply recited financial statement information without analysis or presented boilerplate analyses that did not provide any insight into the companies' past performance or business prospects as understood by management." ("Summary by the Division of Corporation Finance of significant issues addressed in the review of the periodic reports of the Fortune 500 companies," 2/27/2003. <http://www.sec.gov/divisions/corpfin/fortune500rep.htm>)

that managers should provide an *analysis*, explaining management's view of the implications and significance of economic changes at the firm.

2.2 RELATED RESEARCH

Prior research takes three approaches to quantifying disclosure in the MD&A: (1) hand-coded content analysis (Bryan [1997], Rogers and Grant [1997]), (2) survey rankings (Clarkson, Kao, and Richardson [1999], Barron, Kile, and O'Keefe [1999]), and (3) automated text analysis. Under each approach, researchers assess the usefulness of MD&A disclosure by associating their measures of MD&A information or disclosure quality with (1) future operating performance, (2) contemporaneous returns, (3) future returns, and (4) analyst forecast behaviors. The hand-coding and survey approaches result in limited sample size and difficulties for future replication. Because Cole and Jones [2005] and Feldman et al. [2010] provide excellent reviews of the MD&A literature, we highlight a few recent papers that use automated text analysis and discuss how our study advances the literature.

Using automated text analysis, Li [2008] finds that the annual reports (including the MD&A section) of firms with lower earnings and those with positive but less-persistent earnings are more difficult to read. In a follow-up study, Li [2010] finds that firms strategically use the tone of forward-looking statements in the MD&A and that this tone can be used to predict future earnings.

Nelson and Pritchard [2007], Feldman et al. [2010], and Muslu et al. [2009] also use automated text analysis. Nelson and Pritchard extract the cautionary language that either invokes the safe harbor under the Private Securities Litigation Reform Act of 1995 or details risk factors generally found in the MD&A and business description sections of 10-K filings. They take every three adjacent words in a sentence as a unit, called a "trigram," and develop a resemblance score that summarizes whether a trigram is present in both the current year and the previous year. They find that firms that are subject to greater litigation risk update the cautionary language to a larger degree from the previous year; after a decrease in litigation risk the previous cautionary language is not removed.⁶ Feldman et al. find that a positive tone in the MD&A is associated with higher contemporaneous and future returns and that a tone that becomes more negative than the previous year is associated with lower contemporaneous returns. Muslu et al. find that firms provide more forward-looking statements in the MD&A when their stock prices poorly reflect future earnings and these disclosures in turn improve the stock price informativeness of future earnings.

⁶Nelson and Pritchard also calculate a resemblance score for the MD&A (excluding the cautionary language) as a control variable. Their sample is small (1,113 firm-year observations) because the cautionary language section has to be manually identified (even though their resemblance calculations are automated). In their table 3, they report an upward trend in the word count of cautionary language, but not in staleness.

Despite the aforementioned studies, an important facet of MD&A disclosure—to what extent is it modified from the previous year?—is largely unexplored. We introduce a measure of document differences from the information retrieval literature. In addition, our measure is a changes measure by design and therefore better captures *new* information disclosed in the MD&A than the levels measures used by most previous text-analysis studies. Moreover, the regulatory and information environments in which firms operate have changed substantially since most of the MD&A studies were conducted (Francis, Schipper, and Vincent [2002], Bushee, Matsumoto, and Miller [2003], Griffin [2003]). Following these changes, the disclosure mix and the way managers use the MD&A to communicate with the capital markets have also likely changed. Meanwhile, following the Internet revolution, market participants have access to more news outlets and faster information dissemination. Whether and how they use MD&A disclosure has also likely changed. Therefore, we examine MD&A modifications and usefulness in the recent decade.

3. *Measuring Year-over-Year MD&A Modifications*

Internet search engines organize documents into similar groups and compare web users' queries to documents in the search provider's database (e.g., the "Find Similar Documents" function in Google). Many of these search engines have traditionally used the Vector Space Model (VSM) described by Salton, Wong, and Yang [1975] (Singhal [2001]). We use this model to compare a firm's current year MD&A to that from the previous year. The VSM represents a document as a vector in an n -dimensional Euclidean space, where n is the number of unique words in all documents in the sample and the value of each vector element is the frequency of a particular word in that document.⁷ The similarity of any two documents is measured by the angle between the two vectors representing the documents: a smaller angle indicates more similar documents.

Suppose the sample has n unique words. We represent two documents each as an n -dimension vector— v_1 for document 1 and v_2 for document 2:

$$v_1 = (w_1, w_2, \dots, w_{n-1}, w_n) \text{ and } v_2 = (\psi_1, \psi_2, \dots, \psi_{n-1}, \psi_n),$$

where w_i and ψ_i are counts of each word $i \in [1, n]$. The similarity score (*Sim*) is defined as follows:

⁷The VSM may be implemented by using phrases. Using each word as a unit is typically preferable to using phrases or other complex text representations (Salton and Buckley [1988], p. 515, Singhal [2001], p. 5).

$$Sim = \cos(\theta) = \frac{v_1 \cdot v_2}{\|v_1\| \|v_2\|} = \frac{v_1 \cdot v_2}{\|v_1\| \|v_2\|},$$

where θ is the angle between v_1 and v_2 , (\cdot) is the dot product operator, $\|v_1\|$ is the vector length of v_1 , and $\|v_2\|$ is the vector length of v_2 .⁸ This score is bounded between 0 and 1 with a higher score indicating more similarity ($\cos\theta = 1$). The difference score is 1 minus the similarity score.

As a simplified example, consider two documents where the number of unique words, n , is 5. Each document is then represented by a vector of five elements with each value being the number of occurrences of the corresponding word in that document. Three possible cases are as follows:

| | Case (i) | Case (ii) | Case (iii) |
|--------------------------------|--|--|--|
| Vector | $v_1 = (1, 3, 1, 2, 1)$ $v_2 = (1, 3, 1, 2, 1)$ | $v_1 = (1, 1, 0, 1, 1)$ $v_2 = (1, 0, 1, 1, 1)$ | $v_1 = (4, 0, 0, 0, 3)$ $v_2 = (0, 2, 2, 1, 0)$ |
| Length of vector | $\ v_1\ = 4, \ v_2\ = 4$ | $\ v_1\ = 2, \ v_2\ = 2$ | $\ v_1\ = 5, \ v_2\ = 3$ |
| Dot product of v_1 and v_2 | 16 | 3 | 0 |
| Similarity score | 1 | 0.75 | 0 |
| Difference score | 0 | 0.25 | 1 |

The two documents are identical in case (i), slightly different in case (ii), and totally different in case (iii).

Various weighting functions have been developed in the information retrieval literature in applying the VSM. The use of word counts described earlier is referred to as the “term frequency” (TF).⁹ A popular weighting refinement is to multiply TF by the logarithm of M/m , where M is the number of all documents in the sample, m is the number of documents in which that particular word appears, and M/m is referred to as the “inverse document frequency” (IDF). As a result, common words are down-weighted.¹⁰ For example, if a word appears in every document, it has zero

⁸The length of a vector is not the same concept as the dimension of the vector or the length of the document. For example, the dimension of v_1 is n , the length of the document represented by v_1 is $(w_1 + w_2 + \dots + w_n)$, and the length of v_1 is $(w_1^2 + w_2^2 + \dots + w_n^2)^{\frac{1}{2}}$. The similarity score formula is also called the “cosine” measure. It is basically the formula for the uncentered correlation coefficient of two vectors.

⁹We stem each word using the Porter stemming algorithm to abstract away from word tense and form so that we can focus on the underlying word choice (Manning and Schütze [1999]). For example, “faith,” “faithful,” and “faithfulness” are all stemmed to “faith.” Stemming is used to reduce the dimension of vectors and thus computing time.

¹⁰The IDF weighting eliminates the need for a stop-word list as implemented in many related studies, because common words will receive a weight of, or very close to, zero. For this reason, if a firm merely increases boilerplate with no changes in meaningful disclosure, the difference score will be largely unaffected. On the other hand, IDF heavily weights words that are unique to a company, treating them as very “informative” words. To address the concern of overweighing such words, we alternatively use TF with Li’s [2010] stop-word list to remove common words and our results are largely unchanged.

weight (because $\log 1$ is 0). In our empirical execution, we choose the TF-IDF weighting function because of its simplicity and popularity (Salton and Buckley [1988]).

We calculate the difference score between a firm's current year MD&A and that for the previous year and refer to it as the "raw difference score," *Rawscore*. Appendix A presents two examples of MD&A disclosure and their raw scores. We cannot directly use *Rawscore* to compare the degree of MD&A modifications across firms because the raw difference score is a decreasing function of document length. The longer a pair of documents, the more probable a word is included in both documents, leading to a lower likelihood that the documents will differ (see appendix B for an analytical proof). Li [2008] reports that a firm's MD&A length is sticky over time, albeit with a small growth rate, so comparing consecutive-year MD&A disclosures across firms involves two long documents for some firms and two short documents for others. The mechanical relationship between *Rawscore* and document length must be removed before the difference scores are compared across firms. We empirically approximate the functional form of the relation between *Rawscore* and document length by a Taylor expansion at 0 and calculate the expected difference score given the document length. Our MD&A modification score is the raw score minus this expected score and is denoted as variable *Score*.¹¹

Three contemporaneous accounting and finance studies measure document similarities or differences. Hoberg and Phillips [2010] examine whether the similarity of product descriptions in 10-K filings is associated with merger and acquisition decisions. They use the VSM after removing words that appear in at least 95% of 10-K filings.¹² Hanley and Hoberg [2010] use a regression implementation of the VSM to separate the standard content from the informative content in an IPO prospectus; the vector weights are a simple word count. Nelson and Pritchard [2007] use the Ferret Copy Detector software to calculate a resemblance score between the current and prior years' cautionary language in 10-K filings. The software was designed to detect plagiarism and considers the order of words in a trigram; therefore, it detects similarity at a different level (e.g., the same three words in different orders are considered different trigrams). The technique identifies the occurrence but not the frequency of a trigram. In our setting, word frequency in a given MD&A document conveys information, but a very high frequency in the whole sample indicates that the word is commonly used and is thus not very indicative of the content of a given document. Our approach accommodates both features.

¹¹All of our test results are similar if we replace *Score* with *Rawscore* and control for document length by including its natural logarithm.

¹²Our results are largely unchanged if we use TF and remove common words following Hoberg and Phillips's approach.

4. *Sample and Descriptive Statistics*

Our sample period is fiscal years 1997–2006.¹³ It starts with 1997 because we require the previous year’s filings to be on EDGAR and 1996 is the first fiscal year for which almost all companies filed the 10-K electronically. We download 10-K filings from EDGAR and extract the MD&A from Item 7 in each filing.¹⁴ As in Li [2008, 2010], before further processing, we remove tables from the MD&A. We then merge the EDGAR data with Compustat using the Central Index Key (CIK) and exclude observations (1) whose MD&A for the previous year is missing, (2) whose total assets or diluted earnings per share (EPS) for the current year or the previous year are missing, (3) whose year-end stock price is below \$1 (to avoid outliers created by small scalars), or (4) that changed the month of fiscal year end during the year. The screening leaves us with 28,142 firm-year observations. This number is comparable to the number of 28,279 firm-year observations for 1994–2004 in Li [2008]. Table 1 presents descriptive statistics for the sample by fiscal year.

We validate our document difference measure by comparing the MD&A disclosure of two firms in the same industry versus two firms not in the same industry. Because of similar business environments, operating conditions, and specialized industry terminology, the raw difference measure from comparing two firms within an industry should be lower than that from firms not in the same industry. In addition, the measure from comparing firms within an industry should decrease when the definition of industry is stricter. We define “industry” in progressively narrower terms by using the eight-digit Global Industry Classification Standard (GICS) code. Digits 1–2 define the sector (the broadest definition of an industry); digits 1–4 define the industry group; digits 1–6 define the industry; and digits 1–8 define the subindustry (the narrowest definition of an industry).

In each year for each industry definition, we compare a firm to every other firm in the industry and calculate a mean raw difference score for that firm. Averaging this mean difference score of all firms in an industry, we get a data point for this industry for a given year. Figure 1 shows the five data points for each year: four from different industry definitions and

¹³We use annual data to avoid seasonality and Compustat’s updates of originally reported quarterly data (Feldman et al. [2010]). Moreover, Griffin (2003) finds a stronger market reaction to 10-Ks than to 10-Qs.

¹⁴Our sample includes the 10-K and 10-K405. Before April 8, 2002, 10-K405 was filed instead of 10-K if insider trading activity was not disclosed in a timely manner. In addition to Item 7, we also extract Item 7a because some companies use Item 7a for market risk disclosure while others include the disclosure in Item 7. The 10-Ks retrieved from EDGAR are free-form text. We use a variety of string searches to extract MD&A, such as “Item 7,” and exclude the extractions led by phrases such as “please refer to Item 7 for more information” and “Item 7... Page 8” (which is likely to be a Table of Contents entry). We are able to successfully extract MD&A documents from 73% of the 10-K filings that are covered by Compustat. Most of the filings from which we could not extract MD&A refer to other sources for the MD&A.

TABLE 1
Sample Means (Medians) by Year

| <i>Year</i> | <i>Obs.</i> | <i>Score</i> | <i>Assets</i> | ΔEPS | <i>BigN</i> | <i>IO</i> | <i>Analyst</i> | <i>Herf</i> | <i>Litig</i> | $ CAR $ |
|-------------|-------------|--------------------|----------------|--------------------|--------------|------------------|----------------|------------------|--------------|------------------|
| 1997 | 2,655 | -0.005 (-0.049) | 1,176 (143) | -0.005 (0.006) | 0.893 (1) | 0.397 (0.357) | 5.1 (3) | 0.062 (0.043) | 0.307 (0) | 0.037 (0.022) |
| 1998 | 2,767 | 0.052 (0.008) | 1,499 (162) | -0.033 (0.003) | 0.880 (1) | 0.408 (0.367) | 5.5 (4) | 0.061 (0.044) | 0.311 (0) | 0.050 (0.031) |
| 1999 | 2,832 | 0.029 (-0.020) | 1,689 (189) | 0.008 (0.006) | 0.863 (1) | 0.399 (0.349) | 5.5 (4) | 0.060 (0.042) | 0.313 (0) | 0.058 (0.038) |
| 2000 | 2,771 | -0.008 (-0.053) | 1,944 (251) | -0.049 (0.003) | 0.858 (1) | 0.425 (0.386) | 5.4 (3) | 0.057 (0.040) | 0.327 (0) | 0.059 (0.034) |
| 2001 | 2,884 | 0.010 (-0.034) | 2,021 (262) | -0.023 (-0.003) | 0.859 (1) | 0.462 (0.434) | 5.5 (4) | 0.058 (0.045) | 0.347 (0) | 0.041 (0.024) |
| 2002 | 2,716 | -0.002 (-0.038) | 2,590 (319) | 0.036 (0.008) | 0.842 (1) | 0.504 (0.502) | 5.9 (4) | 0.062 (0.051) | 0.331 (0) | 0.036 (0.019) |
| 2003 | 2,907 | -0.002 (-0.035) | 2,939 (348) | 0.060 (0.011) | 0.823 (1) | 0.542 (0.556) | 6.5 (4) | 0.063 (0.049) | 0.334 (0) | 0.030 (0.018) |
| 2004 | 2,931 | -0.023 (-0.053) | 3,507 (420) | 0.017 (0.009) | 0.782 (1) | 0.586 (0.624) | 6.7 (5) | 0.063 (0.048) | 0.323 (0) | 0.028 (0.016) |
| 2005 | 2,910 | -0.019 (-0.043) | 4,484 (486) | 0.004 (0.007) | 0.777 (1) | 0.611 (0.659) | 6.6 (5) | 0.065 (0.050) | 0.318 (0) | 0.026 (0.016) |
| 2006 | 2,769 | -0.032 (-0.059) | 5,413 (584) | 0.008 (0.006) | 0.738 (1) | 0.648 (0.716) | 7.1 (5) | 0.076 (0.049) | 0.310 (0) | 0.024 (0.014) |

Year is the fiscal year. *Score* measures the extent to which two documents are different and has been adjusted for document length (i.e., the number of words). A higher score indicates more differences. *Assets* is the total assets at the end of the fiscal year. ΔEPS is the change in diluted EPS, scaled by the stock price at the end of the fiscal year. Both earnings and price are adjusted for stock splits. The observations with stock prices less than 1 are excluded. *BigN* is 1 if the firm is audited by a Big *N* (*N* is 6, 5, or 4) accounting firm or its predecessor and 0 otherwise. *IO* is the percentage of ownership by institutional investors according to the most recent SEC 13f filings before the firm's 10-K filing, collected by Thomson-Reuters. *Analyst* is the number of financial analysts whose earnings estimates for year $t + 1$ are included in the most recent IBES consensus before the 10-K filing. *Herf* is the Herfindahl index, using the 100 firms with the highest sales in the industry. *Litig* is 1 for high litigation risk firms (the four-digit SIC code is 2833-2836, 8731-8734, 3570-3577, 7370-7374, 3600-3674, or 5200-5961) and 0 otherwise. $|CAR|$ is the absolute value of the cumulative market-adjusted returns over the three days beginning with the 10-K filing date. ΔEPS is winsorized at 1% and 99% each year. The sample has 28,142 firm-year observations. For some variables, the number of observations is slightly smaller.

one from comparing 1,000 random pairs regardless of industry.¹⁵ In each year, the difference measure from comparing two random firms is the highest and the measure decreases monotonically when the industry definition becomes stricter. Such patterns are consistent with our expectations, lending support to the idea that our measure captures document differences. The figure also reveals a downward trend in the difference measure for all groups. This trend is either because firms' MD&A disclosures have become more similar to each other or because document length has increased during our sample period.

After confirming the validity of the raw difference measure, we compare each firm's MD&A document with its previous year's and calculate

¹⁵Comparing these five data points in a given year is valid because each data point is the mean raw score of many firms, each being compared with other firms' documents with different MD&A length.

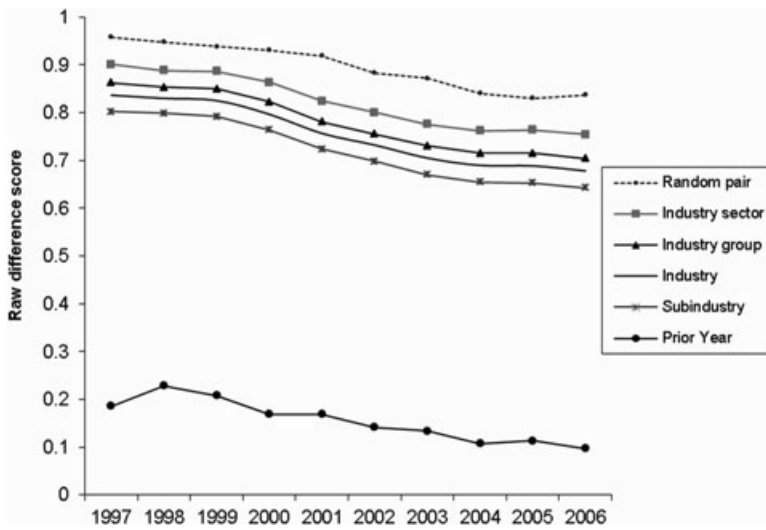


FIG. 1—How different is a firm’s MD&A from other MD&A documents? Raw difference score measures the extent to which two documents are different. A higher score indicates more differences. For each fiscal year, the figure shows the mean raw score of MD&A comparisons (1) of 1,000 randomly selected pairs (the highest line); (2) within an industry sector, industry group, industry, and subindustry; and (3) with the firm’s previous year’s document (the lowest line). A firm is compared with every other firm in the group, the mean *score* of these comparisons are calculated for each firm, and the mean of all firms is the value for that industry definition in a given fiscal year. We use four industry definitions from the GICS code. Digits 1–2 define the sector, digits 1–4 define the industry group, digits 1–6 define the industry, and digits 1–8 define the subindustry. “Sector” is the broadest definition of industry and “subindustry” is the narrowest.

Rawscore. To gauge the magnitude of this score, we additionally plot the average *Rawscore* for each year in figure 1. For example, *Rawscore* is 0.10 for 2006, much lower than the difference scores from comparing firms within the same industry (ranging from 0.64 to 0.75 for the four industry definitions). Of course, a firm’s MD&A document is supposed to be more similar to its own in the previous year (because of the same business characteristics) than to a different firm’s MD&A document. To gauge how high or low a *Rawscore* value is, table B2 of appendix B presents the sample distribution of *Rawscore*. The mean of 0.155 and standard deviation of 0.147 indicate that a *Rawscore* of 0.302 (which is one standard deviation higher than the mean) is higher than 84% of *Rawscore* values in the sample.

Score is calculated in appendix B by adjusting *Rawscore* for document length. Figure 2 presents the distributions of *Rawscore*, *Score*, and MD&A length. The majority of MD&A documents are between 2,000 and 8,000 words; a few are extremely long. As expected, *Rawscore* decreases with document length and *Score* is flat with respect to MD&A length.¹⁶

¹⁶Our sample has 60,296 unique words. The 1st and 99th percentiles of the distribution of document length are 1,107 and 21,703 words. Our results are robust to excluding the 99th percentile of the *Length* observations.

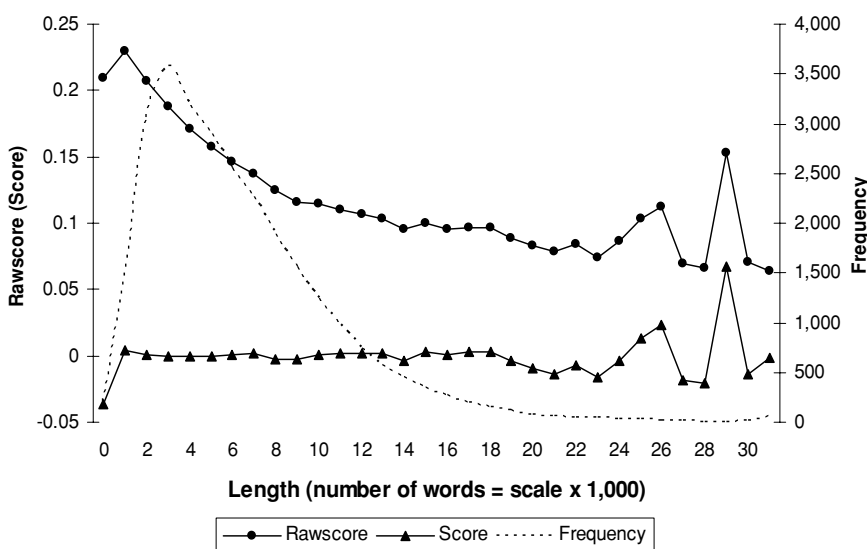


FIG. 2—How different is a firm’s MD&A from that of the previous year? *Rawscore* measures the extent to which two documents are different. A higher score indicates more differences. In theory, *Rawscore* is decreasing in a firm’s MD&A length, which is sticky with an upward drift in practice. We empirically estimate this functional relationship by regressing *Rawscore* on the first five polynomials of *Length*, where *Length* is the number of words in the current year’s MD&A. The fitted score is a firm’s expected raw score given the MD&A length. *Score* is the raw score minus the fitted score. In the graph, *Rawscore* and *Score* are on the left y-axis and *Frequency* of MD&A length is on the right y-axis. This graph shows the distribution of *Length* and the empirical relations between *Rawscore* and *Length* and between *Score* and *Length*. The high volatility on the right-hand side of the graph is due to a small number of extremely long documents.

5. MD&A Modifications After Economic Changes

If managers follow the spirit of MD&A regulations, managers of firms with large economic changes are expected to provide more in-depth discussion and analysis than other firms. On the other hand, managers have substantial discretion in what and how to disclose. Given the costs of preparing long narrative documents, managers may simply use last year’s MD&A as a template and make minor changes. The role of auditors in overseeing MD&A disclosure is limited because they merely review the disclosure for consistency with what is known to them. The risk of an SEC enforcement action is small because of the soft-talk (vs. hard-fact) nature of the disclosure, even though the MD&A is subject to SEC reviews.

5.1 EMPIRICAL MODEL

Economic changes at a firm come from multiple sources, including (1) results of continuing operations, (2) LCR, (3) market risk exposure, and (4) business components (e.g., acquisitions, spin-offs, and discontinued operations). Managers’ view of how and why operating results have changed is

key information required from the MD&A. We measure operating changes by the change in a firm's diluted EPS from the previous year, scaled by the fiscal-year-end stock price, and refer to the variable as ΔEPS .

Also required from the MD&A is managers' view of the firm's liquidity condition and the use and sufficiency of capital resources. We use three measures for liquidity changes: $|\Delta Current|$, $|\Delta Debt\ due|$, and $|\Delta Leverage|$. The current ratio, *Current*, captures the ability to meet working capital needs. The amount of debt due in the coming year, *Debt\ due*, captures the need for short-term liquidity beyond working capital. The amount of total liability, *Leverage*, captures the need for total liquidity. We use free cash flows to capture the sources of cash for capital needs and measure the change in free cash flows as $|\Delta FCF|$. Except for the current ratio, these variables are scaled by total assets before the differences are calculated.

The SEC requires firms to discuss a known market risk that has "affected reported trends or financial condition in the period presented, or is reasonably likely to affect materially future reported results or liquidity."¹⁷ The market risks that companies typically disclose include commodity prices, interest rates, and foreign currency translation. We use a firm's return volatility to proxy for its risk exposure. The greater the exposure, the more volatile the stock price. Return volatility is calculated from weekly raw stock returns to avoid the influence of infrequent and nonsynchronous trading in daily returns. The change in return volatility from the previous year to the current year is denoted as $|\Delta Stdret|$. To mitigate the effect of outliers, all the changes variables are winsorized at the 1st and 99th percentiles before the absolute values are taken.

Another form of economic change is the change in business components. These changes either reflect the execution of a long-term business strategy (e.g., some firms grow by acquisitions while others grow from within) or the cumulative effects of past operating performances reaching a threshold (e.g., some units are discontinued or spun off). We use indicators for substantial asset changes to capture such events: *Acquire* is 1 if total assets increase by at least one-third from the previous year and 0 otherwise; *Downsize* is 1 if total assets decrease by at least one-third from the previous year and 0 otherwise.¹⁸ *Acquire* is 1 for 20.7% of our sample and *Downsize* is 1 for 4.2%.

The MD&A may be modified from the previous year because of the implementation of new MD&A disclosure requirements and accounting standards. We use the fixed-year effects to control for modifications for this reason. We create a dummy variable for each sample year except for 2003. Equation (1) summarizes the empirical model:

¹⁷See SEC staff responses at <http://www.sec.gov/divisions/corpfin/guidance/derivfaq.htm#qgen>.

¹⁸We do not use the SDC data because some firms routinely make small acquisitions. Our method ensures that the changes are major.

$$\begin{aligned}
\text{Score} = & a_0 + a_1|\Delta\text{EPS}| + a_2|\Delta\text{Current}| + a_3|\Delta\text{Debt\&A}| + a_4|\Delta\text{Leverage}| \\
& + a_5|\Delta\text{FCF}| + a_6|\Delta\text{Stdret}| + a_7\text{Acquire} + a_8\text{Downsize} \\
& + \text{year dummies} + e.
\end{aligned}
\tag{1}$$

Next, we hold constant the year-to-year economic changes and examine what firm characteristics, which are typically stable from year to year, are associated with MD&A modifications. These characteristics have been identified by prior research as determinants of corporate voluntary disclosures in other venues. Firm size is expected to be positively associated with the modification score either because large firms are under pressure to modify the MD&A to avoid the political costs of being perceived as being opaque or because small firms might be more likely to cut and paste disclosure from the previous year to save preparation costs. *Size* is the natural logarithm of total assets at the end of the current year. The MD&A is reviewed by auditors, and Big *N* auditors (*N* is 6, 5, or 4 during our sample period) may encourage managers to provide more informative disclosure. *BigN* is 1 if the firm is audited by a Big *N* accounting firm and 0 otherwise.

Institutional investors and analysts are key players in a firm's information environment. Typically, institutional investors and financial analysts demand more corporate disclosures for valuation and monitoring purposes (Bushee and Noe [2000]). This demand may not apply to the MD&A because the disclosure is not a timely source of information. By the time the 10-K is filed, these participants might have already obtained the information elsewhere. We examine institutions' and analysts' associations with MD&A modifications without offering predictions. *IO* is the percentage of ownership by institutional investors according to the most recent SEC 13f filing before the 10-K filing. *Analyst* is the number of analysts whose earnings estimates for the subsequent year are included in the most recent IBES consensus before the 10-K filing.

Informative MD&A disclosure may benefit competitors, so firms in competitive industries may be more concerned about proprietary disclosure costs than those in other industries. We use the Herfindahl index, *Herf*, as a proxy for industry competition and calculate it using the 100 firms (or fewer if the number of firms is less than 100) with the highest sales in the industry (the first six digits of the GICS code are used for industry identifications). A lower *Herf* indicates more intense competition. We predict its coefficient to be positive, meaning that firms in more competitive industries are less likely to make informative MD&A disclosures.

Finally, firms in a more litigious environment may be more likely to provide disclosure to defend themselves against potential lawsuits. Following Francis, Philbrick, and Schipper [1994], we use industry classifications to identify firms that are exposed to high litigation risk, *Litig*. Equation (2) is the expanded empirical model:

TABLE 2
The MD&A Modification Score and Economic Changes

| Panel A: Descriptive Statistics | | | | | | | | |
|--|----------------|--------------------|---------------------|---------------------|----------------|-------------------|----------------|-----------------|
| | Mean | P25 | Median | P75 | | | | |
| <i>Score</i> | 0.000 | -0.090 | -0.041 | 0.047 | | | | |
| <i>EPS</i> | -0.049 | -0.044 | 0.033 | 0.064 | | | | |
| <i>Current</i> | 3.101 | 1.376 | 2.138 | 3.526 | | | | |
| <i>Debt due</i> | 0.020 | 0.000 | 0.003 | 0.021 | | | | |
| <i>Leverage</i> | 0.520 | 0.299 | 0.514 | 0.716 | | | | |
| <i>FCF</i> | -0.025 | -0.060 | 0.018 | 0.071 | | | | |
| <i>Stdret</i> | 0.076 | 0.043 | 0.065 | 0.096 | | | | |
| $ \Delta EPS $ | 0.126 | 0.010 | 0.030 | 0.100 | | | | |
| $ \Delta Current $ | 0.846 | 0.049 | 0.266 | 0.830 | | | | |
| $ \Delta Debt due $ | 0.014 | 0.000 | 0.001 | 0.011 | | | | |
| $ \Delta Leverage $ | 0.071 | 0.014 | 0.038 | 0.088 | | | | |
| $ \Delta FCF $ | 0.083 | 0.014 | 0.044 | 0.104 | | | | |
| $ \Delta Stdret $ | 0.022 | 0.005 | 0.013 | 0.028 | | | | |
| Panel B: Spearman Correlations | | | | | | | | |
| | $ \Delta EPS $ | $ \Delta Current $ | $ \Delta Debt due $ | $ \Delta Leverage $ | $ \Delta FCF $ | $ \Delta Stdret $ | <i>Acquire</i> | <i>Downsize</i> |
| <i>Score</i> | 0.111 | 0.092 | 0.085 | 0.130 | 0.084 | 0.056 | 0.082 | 0.083 |
| $ \Delta EPS $ | | 0.155 | 0.168 | 0.241 | 0.287 | 0.221 | -0.077 | 0.197 |
| $ \Delta Current $ | | | 0.114 | 0.359 | 0.405 | 0.198 | 0.136 | 0.125 |
| $ \Delta Debt due $ | | | | 0.217 | 0.121 | 0.030 | 0.029 | 0.020 |
| $ \Delta Leverage $ | | | | | 0.338 | 0.127 | 0.265 | 0.183 |
| $ \Delta FCF $ | | | | | | 0.218 | 0.096 | 0.188 |
| $ \Delta Stdret $ | | | | | | | 0.015 | 0.097 |
| <i>Acquire</i> | | | | | | | | -0.090 |

See variable definitions in table 3. All the correlations are statistically significant at 5%.

$$\begin{aligned}
 \text{Score} = & a_0 + a_1|\Delta EPS| + a_2|\Delta Current| + a_3|\Delta Debt due| + a_4|\Delta Leverage| \\
 & + a_5|\Delta FCF| + a_6|\Delta Stdret| + a_7Acquire + a_8Downsize + a_9Size \\
 & + a_{10}BigN + a_{11}IO + a_{12}Analyst + a_{13}Herf + a_{14}Litig \\
 & + \text{year dummies} + e.
 \end{aligned} \tag{2}$$

5.2 TEST RESULTS

Panel A of table 2 presents descriptive statistics for the nonindicator variables and panel B reports Spearman correlations. *Score* has reasonable variation, ranging from -0.090 at the 25th percentile to 0.047 at the 75th percentile. *Score* is positively correlated with the absolute changes in EPS, current ratio, short-term debt due, leverage, free cash flows, and return volatility, and with the dummies for the changes in business components.

Table 3 reports the multivariate results, which are robust to heteroskedasticity and within-firm error correlations.¹⁹ The first three columns estimate

¹⁹Serial correlations within a firm and cross-sectional correlations within a sample year may exist in our data. We use Rogers standard errors to address serial correlations and use the

equation (1) in different ways. column 1 uses the raw absolute values of the economic changes. This estimation preserves the magnitude of the changes, but the coefficients cannot be readily compared and the relations between *Score* and the economic changes are restricted to be linear. Column 2 uses the decile rankings with 1 being the smallest absolute change group and 10 being the largest. The coefficients can now be compared; moreover, by using a coarser scale, this approach relaxes the linearity restriction. Because of these advantages, we treat column 2 as our primary analysis. Column 3 aggregates the liquidity and capital resources variables into one variable, *LCR*, which is the mean decile ranking of $|\Delta Current|$, $|\Delta Debtdue|$, $|\Delta Leverage|$, and $|\Delta FCF|$. This approach allows us to compare the degree of disclosure modifications after changes in operations versus changes in *LCR*.

Except for $|\Delta Stdret|$, all the explanatory variables in column 2 have significantly positive coefficients. The magnitude of the coefficients for operations and liquidity variables is between 0.002 and 0.003, indicating that the modification scores from firms in the highest decile are on average 0.018 to 0.027 higher than the scores from those in the lowest decile. This difference is equivalent to about half of the difference between either the 10th and 25th percentiles or between the 25th and 50th percentiles of *Score* and is thus material. If a firm is ranked high for more than one variable, the difference between its modification score and those of other firms would be even larger. These results suggest that managers with large economic changes (except for changes in risk exposure) modify the MD&A more than firms with small economic changes, apparently meeting a minimum disclosure requirement.

In column 3, *LCR* has a coefficient of 0.007 with a *t*-statistic of 13.42, significantly higher than the coefficient of 0.003 on $|\Delta EPS|$ (Wald-test $F = 42.21$). This result suggests that managers modify the MD&A to a larger extent when discussing changes in *LCR* than when discussing operations. Given that firms' earnings announcements, which typically occur a few weeks before the 10-K filing, convey predominately information about operations, the MD&A is perhaps a better information source about *LCR* than about operations.

The last column expands column 2 and examines why firms with similar year-to-year economic changes would modify the MD&A to different degrees. The coefficient on *Size* is significantly positive at 0.007 ($t = 7.61$), suggesting that large firms modify MD&A disclosure from the previous year

fixed-year effects to mitigate the effect of cross-sectional correlations. We do not cluster by years in our main analysis because our panel data are very short in the time series and clustering by 10 years may only add noise to the system (Petersen [2009]). Our results, however, are robust to two-way clustering. We do not include firm-fixed effects in the main analysis because the estimation would examine within-firm variations, whereas our interest is broader in this step and because in equation (2) we specifically examine the firm characteristics that contribute to firm-fixed effects. Our results, however, are robust to adding firm-fixed effects.

TABLE 3
Regression Analysis of the MD&A Modification Score and Economic Changes

| | Measurement Scale of Explanatory Variables | | | |
|--------------------------|--|-----------------------|----------------------|-----------------------|
| | Raw | Deciles | Deciles | Deciles |
| <i>Intercept</i> | -0.026*** (-10.04) | -0.049*** (-14.91) | -0.058 (-15.57) | -0.100*** (-13.36) |
| <i> \Delta EPS </i> | 0.012*** (3.28) | 0.003*** (6.74) | 0.003*** (6.78) | 0.002*** (6.03) |
| <i> \Delta Current </i> | 0.000 (0.36) | 0.002*** (4.30) | | 0.002*** (5.55) |
| <i> \Delta Debt due </i> | 0.202*** (6.30) | 0.002*** (6.45) | | 0.002*** (6.53) |
| <i> \Delta Leverage </i> | 0.128*** (10.10) | 0.003*** (7.37) | | 0.003*** (8.17) |
| <i> \Delta FCF </i> | 0.036*** (3.36) | 0.001** (2.28) | | 0.001*** (3.70) |
| <i>LCR</i> | | | 0.007*** (13.42) | |
| <i> \Delta Stdret </i> | 0.181*** (5.57) | 0.000 (1.20) | 0.000 (0.81) | 0.001* (1.73) |
| <i>Acquire</i> | 0.026*** (9.74) | 0.030*** (11.13) | 0.031*** (11.56) | 0.028*** (10.63) |
| <i>Downsize</i> | 0.026*** (4.46) | 0.037*** (6.68) | 0.037*** (6.62) | 0.037*** (6.55) |
| <i>Size</i> | | | | 0.007*** (7.61) |
| <i>BigN</i> | | | | -0.003 (-0.87) |
| <i>IO</i> | | | | -0.020*** (-4.62) |
| <i>Analyst</i> | | | | -0.001*** (-3.60) |
| <i>Herf</i> | | | | 0.070*** (3.00) |
| <i>Litig</i> | | | | 0.013*** (4.41) |
| <i>Y1997</i> | -0.008** (-2.18) | -0.009** (-2.51) | -0.009*** (-2.61) | -0.009** (-2.33) |
| <i>Y1998</i> | 0.047*** (12.34) | 0.046*** (12.28) | 0.046*** (12.26) | 0.047*** (12.11) |
| <i>Y1999</i> | 0.024*** (6.36) | 0.025*** (6.48) | 0.024*** (6.43) | 0.024*** (6.24) |
| <i>Y2000</i> | -0.013*** (-3.51) | -0.012*** (-3.27) | -0.012*** (-3.30) | -0.013*** (-3.64) |
| <i>Y2001</i> | 0.007** (2.12) | 0.007** (2.21) | 0.007** (2.13) | 0.006* (1.81) |
| <i>Y2002</i> | -0.001 (-0.50) | -0.002 (-0.55) | -0.002 (-0.59) | -0.002 (-0.79) |
| <i>Y2004</i> | -0.020*** (-7.46) | -0.020*** (-7.28) | -0.019*** (-7.25) | -0.019*** (-7.13) |
| <i>Y2005</i> | -0.014*** (-4.91) | -0.013*** (-4.82) | -0.014*** (-4.88) | -0.013*** (-4.63) |

(Continued)

TABLE 3 — Continued

| | Measurement Scale of Explanatory Variables | | | |
|--------------|--|----------------------|-----------------------|----------------------|
| | Raw | Deciles | Deciles | Deciles |
| <i>Y2006</i> | -0.029*** (-10.35) | -0.028*** (-9.99) | -0.028*** (-10.01) | -0.028*** (-9.76) |
| Model F | 73.11*** | 74.74*** | 88.80*** | 60.08*** |
| R^2 | 5.8% | 5.8% | 5.8% | 6.4% |
| Obs. | 28,142 | 28,142 | 28,142 | 28,142 |

Score measures the extent to which two documents are different and has been adjusted for document length. A higher score indicates more differences. ΔEPS is the change in diluted EPS, scaled by the stock price at the end of the fiscal year. *Current* is the current ratio. *Debt due* is the debts due in one year, *Leverage* is the total liabilities, and *FCF* is the free cash flows (i.e., operating cash flows minus capital expenditures) in the current year, all scaled by total assets. *Stdret* is the volatility of weekly returns in the current fiscal year. The change variables are changes from the prior year to the current year and are filled with 0 when the data are missing. The above independent variables are winsorized at 1% and 99% and then the absolute values are calculated. *LCR* is the mean decile ranking of $|\Delta Current|$, $|\Delta Debt due|$, $|\Delta Leverage|$, and $|\Delta FCF|$. *Acquire* is 1 if a firm's asset growth during the year is 1/3 or higher and 0 otherwise. *Downsize* is 1 if the total assets decrease by 1/3 or more during the year and 0 otherwise. *Size* is the natural logarithm of total assets at the end of the fiscal year. See table 1 for the definitions of *BigN*, *IO*, *Analyst*, *Herf*, and *Litig*. The regression estimations are robust to heteroskedasticity and within-firm error correlations. ***, **, and * mark statistical significance at 1%, 5%, and 10% in a two-tailed test, respectively. The *t*-statistics are in the parentheses.

to a greater degree. The coefficients on *IO* and *Analyst* are significantly negative (coefficient = -0.020 and $t = -4.62$ for *IO*; coefficient = -0.001 and $t = -3.60$ for *Analyst*). The positive coefficients of 0.070 on *Herf* ($t = 3.00$) and of 0.013 on *Litig* ($t = 4.41$) suggest that firms facing more competition are less likely to provide informative MD&A disclosure and those in a more litigious environment modify MD&A disclosure to a larger degree from year to year. Except for the coefficients on *IO* and *Analyst*, these findings are consistent with the conclusion of Clarkson et al. [1999] that MD&A disclosure is part of a firm's overall disclosure package, further validating our modification score.

6. Investors' and Analysts' Responses to MD&A Modifications

The usefulness of MD&A disclosure has been a long-standing issue.²⁰ On the one hand, MD&A disclosure was mandated as an important venue for managers to communicate what cannot be delivered by numbers, financial statements, and notes so that investors will understand management's perspective. The modifications are thus expected to be useful. On the other hand, managers may use the financial reporting and disclosure discretion to mislead investors (e.g., Teoh, Welch, and Wong [1998], Bartov and Mohanram [2004], Marquardt and Wiedman [2005]). In particular, Li [2008] concludes that managers use their discretion in preparing narrative disclosure to strategically obfuscate the underlying financial results. If so, even if managers modify the MD&A, users will not have a clear view of the company and therefore their responses to the MD&A may be subdued. Doubts

²⁰Until Griffin [2003], the literature failed to find evidence that investors reacted to 10-K or 10-Q filings.

about MD&A usefulness also arise from its relative lack of timeliness. Even if the MD&A is modified in good faith, the information is likely preempted by other corporate disclosure venues and the media as well as the private information search of sophisticated investors.

6.1 EMPIRICAL MODELS

To examine investors' and analysts' responses to MD&A information, we exclude 3,685 firm-years (13.1% of the sample) that did not announce earnings before the 10-K filing date, because investors and analysts would be reacting to an earnings surprise in addition to new information typically disclosed in the 10-K.²¹ We further exclude 286 firm-years with missing earnings announcement dates because these dates are required in both tests. Investors' responses are measured by $|CAR|$, the absolute value of the cumulative market-adjusted stock return over the three days beginning with the 10-K filing date (Griffin [2003]). For analysts' responses, we first examine the proportion of analysts who revise their earnings forecasts or issue initial forecasts for year $t + 1$ in the 30-day window after the 10-K filing for year t . Then we conduct a very conservative analyst revision test on firm-years that have at least one forecast issued in this postevent window, using only the new forecasts to compute the postevent consensus. $|Revision|$ is the absolute value of the change from the pre- to post-10-K consensus, scaled by the stock price at the end of year t .²² In addition, to put investors' and analysts' reactions to the 10-K in perspective, we compare them to the reactions to earnings announcements and pseudo events. The latter are randomly chosen during fiscal year $t + 1$ after we exclude the earnings announcement and 10-K filing dates for fiscal year t and the earnings announcement dates for the first three quarters of year $t + 1$.

In the multivariate analysis, we control for firm size (*Size*), filing delay (*Filelate*), additional financial information (*NewItems*), and the magnitude of market reaction around the earnings announcement date ($|CAR^{EA}|$). We expect market participants' reactions to large firms' 10-K filings to be smaller than to small firms' because private information search for large firms is more active than for small firms (Grant [1980], Atiase [1985]). *Filelate* is 1 if the filing is more than 90 days after the fiscal year end and 0 otherwise.²³ A late filing indicates a potentially significant problem (Choudhary,

²¹These combined earnings announcement and 10-K filing events are not excluded in Griffin [2003]. Feldman et al. [2010] and You and Zhang [2009] do not exclude these events, either, in documenting a drift in the market reaction after 10-K filings. It is unclear to what extent the drift findings are driven by the postearnings announcement drift of the firms whose earnings news is first released on the 10-K filing date.

²²The test is conservative because if we do not find evidence of analyst reaction for the subsample of revision firms, it would be safe to conclude no analyst reaction for the full sample of revision and no-revision firms.

²³Since 2004, the 10-K filing deadline has been accelerated for certain firms. Firms currently have 60, 75, or 90 days to file the 10-K, depending on company size. If a 10-K is not

Merkley, and Schloetzer [2009]). We do not, however, predict the coefficient on *Filelate* because prior research has documented a price drop when firms fail to file on time and it is unclear whether the market's anticipation of bad news due to the missed deadline was complete.

The MD&A is only part of the 10-K and the literature provides limited guidance on how to quantify and control for other information in the filing. A key portion of such information that investors can obtain uniquely from the 10-K is the additional items reported in the full-fledged financial statements and notes. We proxy for this information using the number of nonmissing and nonzero Compustat items (data1–data399) for the event year and refer to it as *NewItems*. This proxy assumes that firms disclose a similar number of basic financial statement items at the earnings announcement. When this is not the case, the market reaction to firms with more reporting items at the earnings announcement is expected to be larger in magnitude than the reaction to other firms. Our variable $|CAR^{EA}|$ controls for this cross-sectional difference, although its main purpose is to control for the complementary or substitutive relation between the earnings announcement and subsequent 10-K filing. $|CAR^{EA}|$ is the three-day, $[-1, 1]$, cumulative market-adjusted stock return. If the earnings announcement and 10-K filing are complementary, the 10-K of a firm whose earnings announcement is informative is also expected to reveal much new information. If the announcement and filing are substitutes, investors are not expected to react much to the 10-K of a firm whose earnings announcement has already contained significant information. The empirical models are as follows:²⁴

$$|CAR| = b_0 + b_1 Score + b_2 Size + b_3 Filelate + b_4 NewItems + b_5 |CAR^{EA}| + year\ dummies + e \quad (3)$$

$$|Revision| = c_0 + c_1 Score + c_2 Size + c_3 Filelate + c_4 NewItems + c_5 |CAR^{EA}| + year\ dummies + e. \quad (4)$$

6.2 TEST RESULTS

To explore the relative usefulness of 10-K filing events, we plot the stock returns and analyst forecast revisions after 10-K filings along with the responses to earnings announcements and pseudo events. Figure 3 shows that the reaction to earnings announcements is much larger than that to 10-K filings and that the latter is slightly larger than the price movements around pseudo-event dates, suggesting that 10-K filings are informational events but they are less important to investors than earnings announcements.

filed within 90 days after the fiscal year end, it is definitely late under either the old or new deadlines.

²⁴Bryan [1997] controls for only return on assets. Feldman et al. [2010] control for accruals and earnings surprises, but find a significant coefficient only for earnings surprises. We control for the magnitude of earnings announcement return because, arguably, price incorporates a variety of information including return on assets and earnings surprises.

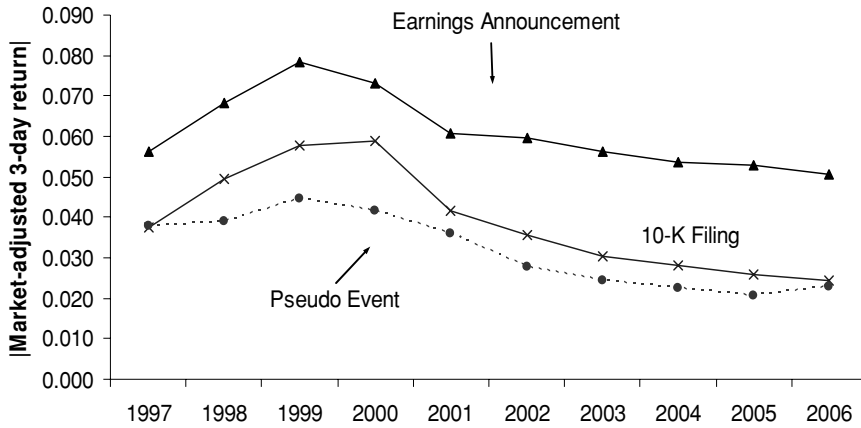


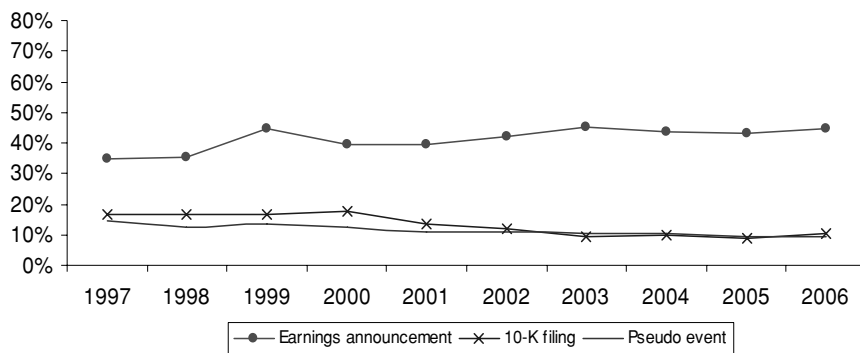
FIG. 3—Price reaction around corporate disclosure events. Price reaction around the fiscal year t earnings announcement is the absolute value of the cumulative market-adjusted return in the event window $[-1, 1]$, where the earnings announcement date is day 0. Price reaction around the 10-K filing for fiscal year t is the absolute value of the cumulative market-adjusted return in the event window $[0, 2]$, where the filing day is day 0. A pseudo-event date is randomly chosen during fiscal year $t + 1$ after earnings announcement and 10-K filing dates are excluded. The price reaction around this date is measured by the absolute value of the cumulative market-adjusted return in $[-1, 1]$, where the pseudo-event date is day 0. Sample means are used for each fiscal year.

For analysts who provide at least one forecast in the $[-90, 30]$ window around the event, figure 4a plots the proportion of analysts who initiate earnings forecasts for year $t + 1$ in the postevent window. Nearly half of the analysts initiate forecasts after earnings announcements, but the percentage after 10-K filings is not distinguishable from that after pseudo events. Figure 4b shows that the proportion of analysts who do *not* revise forecasts is drastically higher after 10-K filings than after earnings announcements. The slight increase after 2002 for 10-K filings might be due to the slight decrease for earnings announcements. That is, earnings announcements have become increasingly informative for earnings revisions, leaving a decreasing amount of information for 10-K events. These patterns suggest that 10-K filings are not important events for analysts in revising earnings estimates.

Tables 4–6 report the response tests, where the analyst revision test uses only firms with at least one forecast in the postevent window. In panel A of table 4, the observations for $|CAR|$ and $|CAR^{EA}|$ are slightly less than 24,171 due to CRSP coverage. At least 9.5% of the sample firms file late. The average number of nonmissing and nonzero Compustat items is 160 and its distribution is fairly symmetric. The mean market reaction at earnings announcements is 6.1%. In panel B, investors' and analysts' responses are significantly positively correlated with the MD&A modification score.

Table 5 reports the multivariate analysis of price reactions, where the estimation is robust to heteroskedasticity and within-firm error correlations.

Panel A: Percentage of Analysts Initiating Earnings Forecasts for Year $t + 1$.



Panel B: Percentage of Analysts Not Revising Earnings Forecasts for Year $t + 1$.

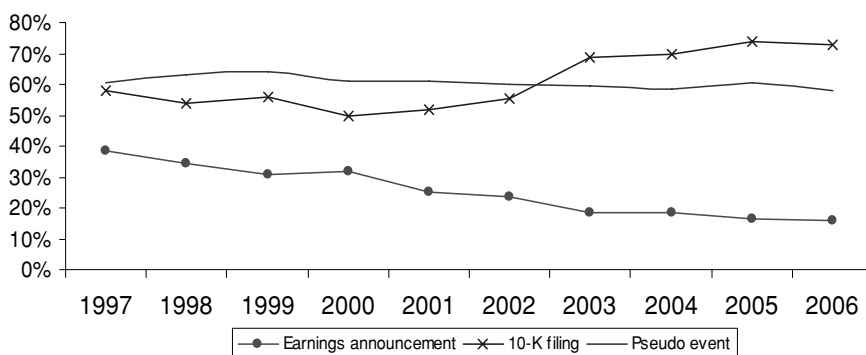


FIG. 4—Analyst responses to corporate disclosure events. The analysts included in the graphs issued at least one forecast in the $[-90, 30]$ window around the event date. A pseudo-event date is randomly chosen during fiscal year $t + 1$ after earnings announcement and 10-K filing dates are excluded.

The left column uses $|CAR|$ as the dependent variable. The intercept is significantly positive because the dependent variable is positive by design, whereas our interest is in the main effect of *Score*. The coefficient on *Score* is significantly positive at 0.016, meaning that if a firm’s score increases from the 25th to 75th percentile of the distribution, the market reaction increases by 0.2%, which is 10% of the median price reaction to 10-K filings. As expected, *Size* is negatively associated with $|CAR|$. *Filelate* is significantly positively associated with $|CAR|$, suggesting that additional information is revealed in the delayed filing. The coefficient on *NewItems* is weakly positive, suggesting that investors respond to the new numerical information in 10-K filings. The coefficient on $|CAR^{EA}|$ is significantly positive, suggesting that earnings announcements complement 10-K filings. The middle column uses the natural logarithm of $|CAR|$ as the dependent variable and the result for *Score* is similar to Column 1, indicating that the least-squares

TABLE 4
Descriptive Statistics for Investors' and Analysts' Responses

| Panel A: Descriptive statistics | | | | | |
|--|--------|--------------------------|--------|--------|-------|
| | Obs. | Mean | P25 | Median | P75 |
| <i>CAR</i> | 23,090 | 3.5% | 0.8% | 2.0% | 4.3% |
| <i>Revision</i> | 10,351 | 0.009 | 0.001 | 0.003 | 0.008 |
| <i>Score</i> | 24,171 | -0.002 | -0.090 | -0.042 | 0.043 |
| <i>Assets</i> | 24,171 | 2,985 | 97 | 339 | 1,247 |
| <i>Filelate</i> | 24,171 | 1 for 9.3% of the sample | | | |
| <i>NewItems</i> | 24,171 | 160 | 143 | 161 | 179 |
| <i>CAR</i> ^{EA} | 23,487 | 6.1% | 1.6% | 3.9% | 8.0% |

| Panel B: Spearman correlations^a | | | | | | |
|---|-----------------|--------------|---------------|-----------------|-----------------|--------------------------|
| | <i>Revision</i> | <i>Score</i> | <i>Size</i> | <i>Filelate</i> | <i>NewItems</i> | <i>CAR</i> ^{EA} |
| <i>CAR</i> | 0.181 | 0.054 | -0.260 | 0.152 | -0.113 | 0.168 |
| <i>Revision</i> | | 0.029 | -0.232 | 0.117 | -0.093 | 0.102 |
| <i>Score</i> | | | -0.010 | 0.047 | 0.113 | 0.056 |
| <i>Size</i> | | | | -0.141 | 0.329 | -0.212 |
| <i>Filelate</i> | | | | | -0.050 | 0.073 |
| <i>NewItems</i> | | | | | | 0.023 |

We exclude 3,685 firm-years that did not announce earnings before the 10-K filing date and 286 firm-years with missing earnings announcement date. |*CAR*| is the absolute value of the cumulative market-adjusted returns over the three days beginning with the 10-K filing date. *Revision* is calculated only for the analysts who revise forecasts within 30 days after the 10-K filing (figure 4b shows that 65% of the analysts do not revise forecasts). *Revision* is the difference between the mean analyst forecasts for year $t + 1$ issued in the 90-day window before the 10-K filing and the mean forecasts issued in the 30-day window after the filing, scaled by the stock price at the end of year t . |*Revision*| is the absolute value of *Revision*. *Score* measures the extent to which two documents are different and has been adjusted for document length. A higher score indicates more differences. *Assets* is the total assets at the end of the fiscal year. *Size* is the natural logarithm of *Assets*. *Filelate* is 1 if the 10-K filing date is at least 90 days after the fiscal year end and 0 otherwise. *NewItems* is the number of nonmissing financial statement items in Compustat. |*CAR*^{EA}| is the absolute value of the three-day, $[-1, 1]$, cumulative market-adjusted returns around the earnings announcement date.

^aThe correlations that are statistically significant at 5% are bolded.

estimation of equation (3) is robust to a violation of the normality assumption (and skewness). In the right column, we subtract the absolute value of the three-day market-adjusted return around a pseudo event from |*CAR*| to control for a firm's normal level of price movement at a typical day. The intercept is not significantly different from 0, indicating that on average the pseudo-event-return adjustment successfully produces an unbiased abnormal return measure. The coefficient on *Score* is 0.007, weakly positive.²⁵ Thus, we conclude that investors appear to use information in MD&A modifications.²⁶

²⁵We do not use this specification as the primary analysis because pseudo-event returns add noise to measurement for individual firms even though the noise aggregates to 0 for the sample as a whole. For example, a pseudo-event date could have an M&A announcement for one firm and be extremely quiet for another firm.

²⁶We identify firms whose 10-K filing dates are clustered and find that the results for this subsample (49.3% of the sample) hold even when we do not adjust returns by the market index. Following Li and Ramesh [2009], we also identify firms whose 10-K filing dates are within five days of the calendar-quarter-end and find robust results for both the calendar-quarter-end clustered (37.3% of the sample) and unclustered subsamples with and without the adjustment by the market index.

TABLE 5
Price Reaction to MD&A Modifications

| | CAR | Log CAR | Adjust CAR by Pseudo-Event Return |
|-------------------|-----------------------|-----------------------|--|
| <i>Intercept</i> | 0.042*** (20.56) | -3.764*** (-60.95) | 0.003 (1.27) |
| <i>Score</i> | 0.016*** (4.80) | 0.217*** (3.34) | 0.007* (1.92) |
| <i>Size</i> | -0.004*** (-19.14) | -0.127*** (-23.34) | -0.001*** (-2.92) |
| <i>Filelate</i> | 0.022*** (10.66) | 0.387*** (12.60) | 0.019*** (8.41) |
| <i>NewItems</i> | 0.000* (1.87) | 0.001*** (3.04) | 0.000 (0.34) |
| CAR ^{EA} | 0.086*** (9.79) | 1.772*** (14.30) | 0.036*** (4.42) |
| <i>Y1997</i> | 0.004*** (4.01) | 0.165*** (4.58) | -0.006*** (4.39) |
| <i>Y1998</i> | 0.016*** (11.56) | 0.426*** (11.24) | 0.005*** (3.34) |
| <i>Y1999</i> | 0.023*** (17.36) | 0.635*** (17.45) | 0.006*** (3.78) |
| <i>Y2000</i> | 0.023*** (15.07) | 0.544*** (14.62) | 0.007*** (4.03) |
| <i>Y2001</i> | 0.007*** (6.09) | 0.235*** (6.27) | -0.002* (-1.64) |
| <i>Y2002</i> | 0.006*** (4.21) | 0.122*** (3.41) | 0.003* (1.91) |
| <i>Y2004</i> | -0.001 (-0.86) | -0.063* (-1.77) | 0.001 (0.75) |
| <i>Y2005</i> | -0.002** (-2.31) | -0.117*** (-3.31) | 0.001 (0.84) |
| <i>Y2006</i> | -0.004*** (-5.72) | -0.223*** (-6.13) | -0.005*** (-4.80) |
| Model F | 140.25*** | 202.66*** | 18.87*** |
| R ² | 12.2% | 11.8% | 2.1% |
| Obs. | 23,083 | 23,083 | 22,594 |

See table 4 for variable definitions. The estimations are robust to heteroskedasticity and within-firm error correlations. For column 3, a pseudo-event date is chosen randomly during fiscal year $t + 1$ after earnings announcement and 10-K filing dates are excluded. The price reaction around this date is measured by the cumulative market-adjusted return in $[-1, 1]$ and the absolute value of this return is subtracted from |CAR|. The decrease of observations from 23,083 to 22,594 is due to no returns data on the selected pseudo-event dates. ***, **, and * mark statistical significance at 1%, 5%, and 10% in a two-tailed test, respectively. The t -statistics are in the parentheses.

Table 6 presents the analyst forecast revision test results, with column 1 using |Revision|, column 2 using $\log|Revision|$, and column 3 using the absolute revision in excess of what analysts would make following a pseudo-event.²⁷ The coefficient on *Score* is statistically insignificant in all the

²⁷We add 0.0001 to every observation before the transformation to retain the 179 firm-year observations that have revision value of 0 (analyst forecasts exist in both the pre- and postevent windows for these observations).

TABLE 6
MD&A Modifications and Analyst Earnings Forecast Revisions

| | $ Revision $ | $Log Revision $ | Adjust $ Revision $ by Revisions Around Pseudo-Event |
|------------------|-----------------------|-----------------------|---|
| <i>Intercept</i> | 0.015** (9.79)* | -4.771** (-47.95)* | 0.010*** (7.51) |
| <i>Score</i> | 0.000 (0.31) | 0.042 (0.44) | -0.000 (-0.24) |
| <i>Size</i> | -0.002** (-11.26)* | -0.138** (-14.97)* | -0.001*** (-9.30) |
| <i>Filelate</i> | 0.005** (5.74)* | 0.368** (7.48)* | 0.003*** (3.13) |
| <i>NewItems</i> | 0.000 (1.29) | 0.001 (1.39) | -0.000 (-0.20) |
| $ CAR^{EA} $ | 0.018** (4.861)* | 1.134** (6.01)* | 0.002 (0.42) |
| <i>Y1997</i> | 0.000 (0.47) | -0.029 (-0.615) | 0.000 (0.57) |
| <i>Y1998</i> | 0.002** (3.08)* | 0.105** (2.25) | 0.001* (1.78) |
| <i>Y1999</i> | 0.001 (1.00) | 0.064 (1.36) | 0.000 (0.63) |
| <i>Y2000</i> | 0.005** (6.57)* | 0.295** (6.03)* | 0.003*** (3.51) |
| <i>Y2001</i> | 0.002** (3.45)* | 0.143** (3.52)* | 0.002*** (2.66) |
| <i>Y2002</i> | 0.005** (7.42)* | 0.318** (7.71)* | 0.003*** (3.83) |
| <i>Y2004</i> | 0.001 (0.99) | -0.054 (-1.43) | 0.001 (0.95) |
| <i>Y2005</i> | 0.001 (1.15) | 0.015 (0.40) | 0.000 (0.54) |
| <i>Y2006</i> | 0.001 (1.57) | 0.015 (0.37) | 0.000 (0.33) |
| Model F | 26.51*** | 44.57*** | 11.56*** |
| R^2 | 5.6% | 8.7% | 2.3% |
| Obs. | 10,165 | 10,165 | 9,919 |

See table 4 for variable definitions. The estimations are robust to heteroskedasticity and within-firm error correlations. For column 3, a pseudo-event date is chosen randomly during fiscal year $t + 1$ after earnings announcement and 10-K filing dates are excluded. The absolute value of consensus analyst forecast revision from the 90-day window before to the 30-day window after this date is subtracted from $|Revision|$. The decrease of observations from 10,165 to 9,919 is due to no revision data around the selected pseudo-event dates. ***, **, and * mark statistical significance at 1%, 5%, and 10% in a two-tailed test, respectively. The t -statistics are in the parentheses.

estimations, even though the results for the control variables are largely consistent with table 5. We conclude that analysts do not use the MD&A in revising earnings forecasts.²⁸ The finding that investors respond to the MD&A but that analysts do not is consistent with our finding in section 5 that managers apparently modify the MD&A more after large changes in

²⁸It is possible that analysts do not revise a forecast unless the revision is at least a penny. If so, our forecast revision variable has captured larger revisions and missed small revisions. We find that even those large revisions are not associated with the MD&A modification score.

LCR than after large changes in operations. The LCR information is useful for predicting future cash flows, but less useful for analysts' year-ahead earnings predictions.

7. *Hand-Coding MD&A Documents*

To further validate our VSM-based modification score and probe what types of information are modified, we hand-code a subset of MD&A documents. We sort the sample into quintiles (each year) by the MD&A modification score and randomly select 50 firms from the highest quintile and 50 firms from the lowest quintile. Each firm's MD&A documents for the current year (t) and the previous year ($t - 1$) are compared and coded.²⁹ On average, an MD&A document is 17 pages long.

We first identify the sections of MD&A disclosure. Almost all firms clearly label the sections (except for "overview" and "cautionary language") but section names vary from firm to firm and some firms use several sections to discuss what would be contained in one section by another firm. We organize the reported sections into "overview," "operations," "LCR," "accounting policies and estimates," "risk factors," "cautionary language," "recent accounting standards," and "other."³⁰

Within each section, we read through the paragraphs and sentences and identify the basic units of discussion, referred to as *aspects*. For example, in the "operations" section, the discussion of current-year sales and its change from the previous year is one aspect and the discussion of profit margin is another aspect. Firms typically discuss one aspect in a paragraph, although an aspect is occasionally discussed in more than one paragraph. For each aspect we code whether it appears in both year t and year $t - 1$ ("same aspect") or not ("different aspect") and, if it is a "same aspect," whether the discussion in year t has been substantially changed from year $t - 1$. If a firm merely changes numbers, a few noncrucial phrases, or one or two noncrucial sentences in a long paragraph, the change is not considered substantial. In addition, for each aspect appearing in year t , we code whether the discussion is interpretive and forward-looking. If the discussion provides explanations, analysis, or a great amount of detail, it is "interpretive." If the discussion includes specific strategies, plans, projections, and managers' assessment of the input factors and market/industry trends, it is "forward-looking."

Panel A of table 7 presents the mean and median hand-coded measures separately for the high- and low-score groups along with the VSM-based

²⁹Two coders with an accounting background are trained in pilot coding and receive an equal number of high- and low-score firms. Both the coding and review are blind to a firm's high- vs. low-score status.

³⁰The "other" section includes "discontinued operations," "subsequent events," "Y2K," "off-balance sheet arrangements," "related party transaction," "trends & management actions," "outlook," etc.

modification score. The high-score group discusses a similar number of aspects as the low-score group (35.1 vs. 31.3, $t = 1.57$), but uses significantly more words (7,424 vs. 5,214, $t = 2.88$). The percentage of different aspects for the former is significantly higher than for the latter (42.8% vs. 17.4%, $t = 7.07$), indicating that high-score firms add new aspects and remove old aspects to a larger extent than the low-score firms. Even for “same aspect,” the discussion by the high-score group is revised to a larger degree than by the low-score group (53.2% vs. 27.4%, $t = 7.02$).³¹ These results further validate our VSM-based modification score.

Regarding the types of information firms disclose, the last two rows of panel A indicate that the two groups do not differ much in the quantity of interpretive disclosures (mean of 41.8% vs. 34.4%, $t = 1.79$; median of 40.3% vs. 33.7%, $Z = 1.18$), but the high-score group provides significantly more forward-looking disclosure than the low-score group (7.2% vs. 4.2%, $t = 2.51$). Panel B presents key measures by section. The discussion of operating results accounts for 36% of the words used in the MD&A (untabulated) and includes about 11 aspects. The discussion of LCR accounts for 22% of the words and covers about eight aspects. For both sections, the high-score group includes different aspects and modifies the discussion of “same aspect” from the previous year to a larger extent than the low-score group. Furthermore, we observe that managers discuss more different aspects about LCR than about operations (23.2% vs. 14.9%, $t = 2.49$), but modify the discussion of “same aspect” to a larger degree for operations than for LCR (44.0% vs. 54.1%, $t = -2.43$).

Not all firms provide disclosure beyond operations and LCR. When they do, we observe higher modifications by the high-score group in “overview,” “accounting policies and estimates” (through modifying the discussion of existing aspects), “risk factors,” and “other” (through modifying the discussion of existing aspects). Yet, the statistical significance is the strongest for operations and LCR and the degree of modifications is greater for these sections than for risk disclosure and cautionary language.

We also probe the extent to which the MD&A discussions about LCR are *not* preempted in the earnings announcement press releases and are *not* duplicated in the notes of the same 10-K filing. That is, we ask how new the LCR information is. We randomly select 20 firms in the high-score group and trace each aspect coded in the MD&A to the earnings announcement release. We observe that 98.4% of the MD&A LCR discussions either do not appear in the release or are much more thorough than that in the release, while 44.4% of the firms do not mention any LCR issue at all in the release. Managers are not supposed to merely repeat in the MD&A what is already disclosed in the notes. We trace the MD&A LCR discussions to various notes

³¹In untabulated analysis, we sum the percentage of different aspects and the percentage of different discussion of “same aspect” for each section and weight the sum by the number of words in each section. This measure is highly correlated with the VSM-based modification score (Spearman correlation coefficient = 0.671).

and observe that 61.2% of the MD&A discussion provides details and interpretation beyond what is available in the notes. It appears that much *new* LCR information is provided by firms that modify the MD&A diligently. Such information could be helpful to investors in projecting future cash flows, while not as useful to analysts in revising year-end earnings estimates.

In sum, the analysis of the hand-coded sample bolsters our case that the VSM-based modification score captures changes in narrative disclosure. Moreover, the analysis shows that many of the MD&A changes are related to operations and, in particular, LCR.

8. MD&A Modifications and Usefulness over Time

The previous two sections use data from a 10-year period and make inferences about MD&A modification and usefulness for the decade as a whole. The past decade has seen a tightening of regulations. The Sarbanes-Oxley Act was passed in July 2002 and most of the recommended changes took place in 2003. These regulations require firms to provide additional disclosure, for example, on internal control and off-balance-sheet arrangements. After 2003, firms are also explicitly required to discuss critical accounting

TABLE 7
Hand-Coded MD&A Subsample

| Panel A: Firm-Level Analysis | | | | | | | |
|--|-----------------|------------------|--------|-----------------|--------|----------|----------|
| | Year | High-Score Firms | | Low-Score Firms | | T-test | Wilcoxon |
| | | Mean | Median | Mean | Median | <i>t</i> | Z |
| <i>Score</i> | <i>t - 1, t</i> | 0.229 | 0.182 | -0.129 | -0.128 | 15.50*** | 8.61*** |
| <i>Rauscore</i> | <i>t - 1, t</i> | 0.382 | 0.351 | 0.041 | 0.036 | 13.26*** | 8.61*** |
| <i>Words</i> | <i>t</i> | 7,424 | 6,744 | 5,214 | 5,182 | 2.88*** | 2.31** |
| <i>Number of aspects</i> | <i>t</i> | 35.1 | 31.5 | 31.3 | 30.5 | 1.57 | 1.19 |
| <i>Different aspect (%)</i> | <i>t - 1, t</i> | 42.8% | 40.8% | 17.4% | 15.4% | 7.07*** | 6.21*** |
| <i>Same aspect (%)</i> | <i>t - 1, t</i> | 71.6% | 71.2% | 86.5% | 88.8% | -6.57*** | -5.67*** |
| <i>Change in discussion of same aspect (%)</i> | <i>t - 1, t</i> | 53.2% | 52.6% | 27.4% | 22.3% | 7.02*** | 5.79*** |
| <i>Interpretive (%)</i> | <i>t</i> | 41.8% | 40.3% | 34.4% | 33.7% | 1.79* | 1.18 |
| <i>Forward looking (%)</i> | <i>t</i> | 7.2% | 4.1% | 4.2% | 3.6% | 2.51** | 1.85* |

1. Sample firms are sorted into quintiles each year by the MD&A modification score. Fifty firms are randomly selected from the highest (lowest) quintile and are referred to as the "high (low)-score group." Each firm's MD&A documents for the current year (*t*) and the previous year (*t - 1*) are compared and coded.

2. "Aspect" is the basic unit of discussion, for example, sales performance in the current year. If an aspect appears in both years, it is a "same aspect"; otherwise, it is a "different aspect." Both rows are presented as percentages, using the number of current-year aspects as the scalar.

3. For a "same aspect," the discussions in both years are compared and the aspect receives a "change" code if the change is substantial. This row is presented as percentages, using the number of "same aspect" as the scalar.

4. The last two rows reflect the type of discussion for an aspect. If the discussion provides explanations, analysis, and a great amount of detail, the aspect is coded as "interpretive." If the discussion includes statements regarding a firm's strategy, plans, projections, and managers' assessment of the input factors and market/industry trends, the aspect is coded as "forward-looking." Each row is presented as percentages, using the number of current-year aspects as the scalar.

5. ***, **, and * mark statistical significance at 1%, 5%, and 10% in a two-tailed test, respectively.

TABLE 7 — *Continued*

| Section | High-Score Firms | | Low-Score Firms | | T-test | Wilcoxon |
|-------------------------------------|------------------|--------|-----------------|--------|----------|----------|
| | Mean | Median | Mean | Median | <i>t</i> | <i>Z</i> |
| Overview: | 38 firms | | 38 firms | | | |
| Words | 755 | 535 | 555 | 455 | 1.40 | 0.84 |
| Number of aspects | 3.7 | 2 | 3.5 | 3 | 0.22 | -0.57 |
| Different aspect | 42.8% | 36.7% | 12.6% | 0.0% | 3.68*** | 2.99*** |
| Change in discussion of same aspect | 54.9% | 50.0% | 19.3% | 0.0% | 3.77*** | 3.52*** |
| Operations: | 50 firms | | 50 firms | | | |
| Words | 2,601 | 2,522 | 1,919 | 1,833 | 2.39** | 1.71* |
| Number of aspects | 11.5 | 11 | 10.8 | 10 | 0.73 | 0.88 |
| Different aspect | 22.5% | 16.7% | 7.2% | 0.0% | 4.15*** | 4.77*** |
| Change in discussion of same aspect | 69.3% | 76.0% | 38.9% | 40.0% | 6.64*** | 5.60*** |
| Liquidity and capital resources: | 50 firms | | 50 firms | | | |
| Words | 1,700 | 1,331 | 1,121 | 969 | 1.92* | 1.87* |
| Number of aspects | 8.0 | 7 | 7.4 | 7 | 0.92 | 0.54 |
| Different aspect | 34.1% | 21.1% | 12.2% | 11.1% | 4.49*** | 3.67*** |
| Change in discussion of same aspect | 56.5% | 60.0% | 31.5% | 28.6% | 4.32*** | 4.06*** |
| Accounting policies and estimates: | 29 firms | | 31 firms | | | |
| Words | 1,364 | 1,228 | 770 | 702 | 3.94*** | 3.51*** |
| Number of aspects | 5.7 | 6 | 4.6 | 4 | 2.14** | 1.93* |
| Different aspect | 38.7% | 22.2% | 28.1% | 0.0% | 0.94 | 1.66* |
| Change in discussion of same aspect | 27.3% | 25.0% | 3.4% | 0.0% | 4.03*** | 4.01*** |
| Risk factors: | 39 firms | | 39 firms | | | |
| Words | 1,393 | 538 | 837 | 419 | 1.74* | 1.40 |
| Number of aspects | 4.8 | 4 | 4.8 | 3 | 1.43 | 1.73* |
| Different aspect | 72.3% | 25.0% | 15.3% | 0.0% | 1.44 | 2.43** |
| Change in discussion of same aspect | 31.5% | 2.5% | 11.0% | 0.0% | 2.44** | 2.52** |
| Cautionary language: | 34 firms | | 36 firms | | | |
| Words | 193 | 149 | 233 | 225 | 1.00 | -1.92* |
| Number of aspects | 1.2 | 1 | 1.1 | 1 | 1.38 | 1.28 |
| Different aspect | 22.1% | 0.0% | 8.3% | 0.0% | 1.62 | 1.69* |
| Change in discussion of same aspect | 14.8% | 0.0% | 6.1% | 0.0% | 1.07 | 1.10 |
| Recent accounting standards: | 32 firms | | 23 firms | | | |
| Words | 404 | 442 | 476 | 381 | -0.82 | 0.23 |
| Number of aspects | 3.1 | 3 | 3.4 | 3 | -0.52 | 0.28 |
| Other: | 30 firms | | 25 firms | | | |
| Words | 470 | 449 | 385 | 293 | 1.03 | 1.01 |
| Number of aspects | 2.3 | 2 | 2.6 | 1 | 0.44 | 0.87 |
| Different aspect | 47.1% | 33.3% | 30.9% | 0.0% | 1.14 | 1.06 |
| Change in discussion of same aspect | 49.3% | 50.0% | 16.7% | 0.0% | 2.73*** | 2.31*** |

If a firm does not have a given section, it is excluded from comparisons for that section.

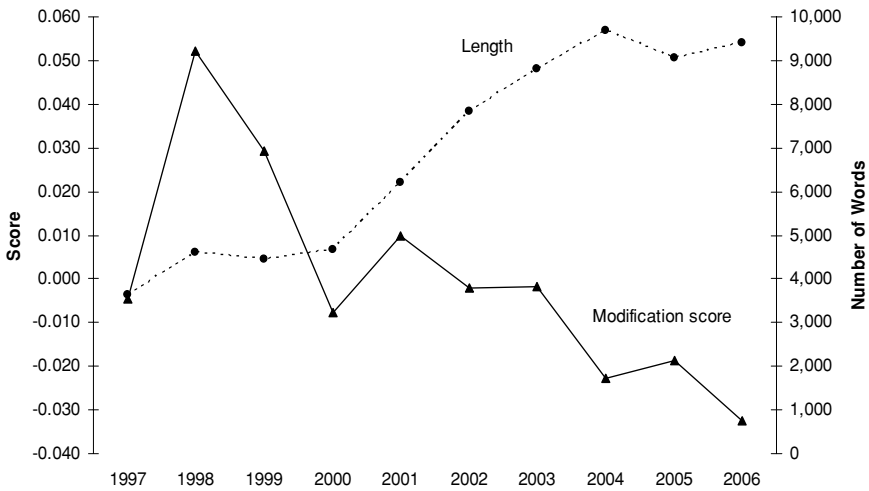


FIG. 5—Distributions of MD&A modification scores and document length. *Score* measures the extent to which two documents are different and has been adjusted for document length (i.e., the number of words). A higher score indicates a larger difference. *Score* uses the left *y*-axis and *Length* uses the right *y*-axis. Sample means are used for each fiscal year.

estimates in the MD&A. Specific regulatory requirements elicit disclosure that managers may otherwise not provide, increasing the degree of MD&A modifications in the initiation year, but potentially leading to meaningless boilerplate discussion merely out of compliance in subsequent years. It is exactly for this reason that the SEC has adopted the management approach, leaving the disclosure decisions to managers. This tradeoff is also why the SEC's 2003 MD&A guidance specifically reminded firms to provide *meaningful* discussion and analysis. Whether the increased regulation in the past decade has led to more or less informative MD&A disclosure is an empirical question of relevance to investors and regulators.

We show in figure 5 the distributions of sample mean *Score* and *Length* each year in the sample period. MD&A length increases over time, possibly triggered by more regulations as well as a more litigious environment after the Enron and WorldCom scandals. In contrast, *Score* generally declines over these years. The dramatic increase in *Score* in 1998 might be due to the SEC's push for "plain English" (SEC [1998], Loughran and McDonald [2008]) or triggered by Y2K disclosures. *Score* for fiscal year 2001 is relatively high, possibly due to the Enron scandal (Leuz and Schrand [2009]) and disclosure by airlines and insurance carriers of the effect of terrorists' attacks (Carter and Simkins [2004]).³² Other than these two years, the overall trend in *Score* is downward, which is confirmed by the year-fixed

³²Using page counts, Leuz and Schrand [2009] find that firms responded to the exogenous shock of a desire for transparency after the Enron collapse with longer 10-K filings.

TABLE 8
Price Reaction to MD&A Modifications over Time

| | CAR | | Log CAR | | Adjust CAR by Pseudo-Event Return | |
|-----------------------|--------|---------------------|---------|---------------------|--|---------------------|
| | Coeff. | <i>t</i> -Statistic | Coeff. | <i>t</i> -Statistic | Coeff. | <i>t</i> -Statistic |
| <i>Intercept</i> | 0.041 | 20.49*** | -3.774 | -61.06*** | 0.003 | 1.16 |
| <i>Score</i> | 0.002 | 0.50 | 0.129 | 0.65 | -0.000 | -0.07 |
| <i>Y1997</i> | 0.004 | 3.99*** | 0.166 | 4.57*** | -0.006 | -4.30*** |
| <i>Y1998</i> | 0.015 | 11.23*** | 0.408 | 10.49*** | 0.005 | 2.99*** |
| <i>Y1999</i> | 0.023 | 17.36*** | 0.632 | 17.25*** | 0.006 | 4.03*** |
| <i>Y2000</i> | 0.023 | 14.89*** | 0.546 | 14.60*** | 0.007 | 4.11*** |
| <i>Y2001</i> | 0.008 | 6.22*** | 0.236 | 6.28*** | -0.002 | -1.58 |
| <i>Y2002</i> | 0.006 | 4.18*** | 0.120 | 3.35*** | 0.003 | 1.92* |
| <i>Y2004</i> | -0.001 | -1.48 | -0.073 | -2.00** | 0.000 | 0.38 |
| <i>Y2005</i> | -0.002 | -2.76*** | -0.118 | -3.32*** | 0.001 | 0.54 |
| <i>Y2006</i> | -0.005 | -5.58*** | -0.240 | -6.21*** | -0.005 | -4.51*** |
| <i>Score * Y1997</i> | 0.006 | 0.72 | 0.005 | 0.02 | 0.009 | 0.83 |
| <i>Score * Y1998</i> | 0.026 | 2.71*** | 0.478 | 1.85* | 0.021 | 1.84* |
| <i>Score * Y1999</i> | 0.018 | 2.07** | 0.287 | 1.16 | -0.005 | -0.46 |
| <i>Score * Y2000</i> | 0.041 | 2.57*** | 0.220 | 0.83 | 0.030 | 1.72* |
| <i>Score * Y2001</i> | 0.005 | 0.52 | 0.059 | 0.22 | 0.002 | 0.16 |
| <i>Score * Y2002</i> | 0.005 | 0.37 | -0.450 | -1.55 | 0.006 | 0.42 |
| <i>Score * Y2004</i> | -0.007 | -0.99 | -0.281 | -0.90 | -0.007 | -0.82 |
| <i>Score * Y2005</i> | -0.007 | -1.01 | 0.043 | 0.14 | -0.008 | -0.94 |
| <i>Score * Y2006</i> | 0.000 | 0.03 | -0.396 | -1.05 | 0.001 | 0.11 |
| <i>Size</i> | -0.004 | -19.29*** | -0.127 | -23.35*** | -0.001 | -3.00*** |
| <i>Filelate</i> | 0.022 | 10.67*** | 0.399 | 12.61*** | 0.018 | 8.40*** |
| <i>NewItems</i> | 0.000 | 2.29** | 0.001 | 3.22*** | 0.000 | 0.56 |
| CAR ^{EA} | 0.086 | 9.74*** | 1.770 | 14.28*** | 0.036 | 4.40*** |
| Model F | | 87.85*** | | 124.87*** | | 12.27*** |
| <i>R</i> ² | | 12.3% | | 11.9% | | 2.2% |
| Obs. | | 23,083 | | 23,083 | | 22,594 |

See variable definitions in table 4. The estimations are robust to heteroskedasticity and within-firm error correlations. With |CAR| being the dependent variable, the *F*-statistics from the Wald tests of (*Score* + *Score * Y1998*), (*Score* + *Score * Y1999*), and (*Score* + *Score * Y2000*) are 12.04, 8.28, and 8.28, respectively. The *F*-statistics are 13.98, 7.56, and 3.66 when Log|CAR| is the dependent variable and are 4.67, 0.37, and 3.24 when |CAR| is adjusted by the pseudo-event return. ***, **, and * mark statistical significance at 1%, 5%, and 10% in a two-tailed test, respectively. The *t*-statistics are in the parentheses.

effects in table 3. The SEC's 2003 guidance did not reverse this trend. Over time the MD&A has become longer but looks more like what investors saw in the previous year.

The market reaction to MD&A information might have changed over time and whether the MD&A has been useful in recent years is an empirical question. On the one hand, Griffin (2003) reports an increase in market reaction to 10-K filings over his sample period of 1996–2001. Market reaction after 2004 has likely increased following the accelerated 10-K filing deadlines, because the acceleration should improve the timeliness of MD&A disclosure. On the other hand, there are reasons to believe that investors' reaction to MD&A disclosure has weakened. The past decade has seen increased corporate interim disclosures (e.g., expanded earnings

announcements, increased use of quarterly conference calls, and increased scope of 8-K filings), more media outlets, faster information dissemination, and increased ease of private information search. These factors make it likely that MD&A information is increasingly preempted by other information sources over time.

Figure 3 shows that the price reaction to 10-K filings increased for fiscal years 1997 to 2000 and was as high as about 6% for 2000, consistent with Griffin [2003], who uses calendar filing years whereas we use fiscal years. The price reaction, however, has declined since 2000 to as low as 2% in 2006. This decline is confirmed by the year dummies in table 8. More importantly, tests of the sum of coefficients on *Score* and the interaction of *Score* with the year dummies indicate that the association of price reaction and MD&A modification scores is significantly positive in 1998–2000, but disappears after 2000. This result is confirmed in yearly estimations, where we allow all the coefficients to vary each year (untabulated).

9. Conclusion

Our study introduces a document modification score from the information retrieval literature to detect year-over-year changes in the extensiveness of a firm's MD&A disclosure. A document that is very similar to that from the previous year does not reveal much new information. We find that firms with larger economic changes modify the MD&A more than those with smaller economic changes, suggesting that firms meet a minimum MD&A disclosure requirement. In addition, managers modify the MD&A to a larger extent to reveal information about the firm's LCR than about operations. Consistent with this result, we find that equity investors react to the MD&A, but that analysts do not revise their earnings forecasts, suggesting that the MD&A contains new information for investors to predict cash flows in future periods even though it provides little new information about the subsequent period's earnings.

Despite firms apparently meeting a minimum MD&A requirement and investors reacting to MD&A information, the overall trend in MD&A modification scores is downward over the past 10 years. This trend is coupled with an increase in MD&A length, suggesting that managers have increased the use of boilerplate disclosure. Moreover, we find that the price reaction to 10-K filings has also declined over time and that it is not associated with MD&A modifications after 2000, suggesting declining or insignificant MD&A usefulness in recent years.

Overall, we introduce a measure for narrative disclosure that could be applied in a variety of other accounting research settings. We document an important facet of MD&A disclosure—modifications. Our findings are likely to be useful to investors and regulators in reevaluating the MD&A regulation and disclosure behaviors.

APPENDIX A. EXAMPLES OF MD&A DISCLOSURE ABOUT CAPITAL EXPENDITURES

Example 1: Brunswick Corp

| Fiscal Year 2002 10-K | Fiscal Year 2003 10-K |
|---|---|
| <p>The Company invested \$112.6 million, \$111.4 million and \$156.0 million in capital expenditures in 2002, 2001, and 2000, respectively. The largest portion of these expenditures was made for on-going investments to introduce new products, expand product lines and achieve improved production efficiencies and product quality.</p> | <p>The Company invested \$159.8 million in capital expenditures in 2003. The largest portion of these expenditures was made for investments to introduce new products and expand product lines in the Marine Engine, Fitness and Bowling & Billiards segments, and achieve improved production efficiencies and product quality. The most significant expenditures in 2003 relate to the equipment needed for production of Verado, the Marine Engine segment's new series of high-horsepower outboard engines introduced in 2004, and the conversion of 13 bowling centers to Brunswick Zones. The Company anticipates spending approximately \$180.0 million for capital expenditures in 2004. About one-half of the capital spending covers investments in new and upgraded products, and plant capacity expansion in the Marine Engine and Boat segments, about one-third for profit maintaining capital and the balance targeted toward cost reductions and investments in information technology.</p> |

Rawscore = 0.86.

Example 2: Interline Brands Inc.

| Fiscal Year 2005 10-K | Fiscal Year 2006 10-K |
|--|---|
| <p>Capital expenditures were \$7.9 million in 2005 as compared to \$6.8 million in 2004. Capital expenditures as a percentage of sales were 0.9% in year 2005 and 0.9% in 2004.</p> <p>Acquisition expenditures were \$73.2 million in 2005, which consisted of \$69.5 million related to our Copperfield acquisition in July 2005 and \$3.6 million related to our Florida Lighting acquisition in November 2003, compared to \$0.6 million in 2004, which related to our Florida Lighting acquisition.</p> | <p>Capital expenditures were \$7.8 million in 2006, \$7.9 million in 2005 and \$6.8 million in 2004. Capital expenditures as a percentage of sales were 0.7%, 0.9% and 0.9% in 2006, 2005, and 2004, respectively.</p> <p>Acquisition expenditures were \$131.5 million in 2006, which consisted primarily of 130.8 million related to our AmSan acquisition in July 2006, \$73.2 million in 2005, which consisted of \$69.5 million related to our Copperfield acquisition in July 2005 and \$3.6 million related to our Florida Lighting acquisition in November 2003 and \$0.6 million in 2004, which related to our Florida Lighting acquisition.</p> |

Rawscore = 0.17.

The raw scores from comparing the *excerpts* of MD&A documents are not comparable to the raw scores from comparing the whole MD&A documents.

APPENDIX B. ADJUSTING THE RAW DIFFERENCE SCORE BY DOCUMENT LENGTH

1. The relation between the raw difference score and document length:

Let us assume a simple scenario, where (1) a manager must choose from a population of n words ($w_1, w_2, w_3, \dots w_n$) to write a document, (2) the firm's document has r number of words, and (3) each word is used in a document only once.

| Previous year's document | | | | | Current year's document | | | | |
|--------------------------|-------|-----|-----------|-------|-------------------------|---|---|---|---|
| w_1 | w_2 | ... | w_{r-1} | w_r | ? | ? | ? | ? | ? |

Question 1: How many ways can the manager select words for the current year's document?

Answer: In statistical terms, we ask how many ways one can combine r objects chosen from n objects without replacement:

$${}_n C_r = \frac{n!}{r!(n-r)!}$$

Question 2: How many ways can the manager select words for the current year's document without repeating the words used in the previous year?

Answer: In statistical terms, we ask how many ways one can combine r objects chosen from $(n - r)$ objects without replacement:

$${}_{n-r} C_r = \frac{(n-r)!}{r!(n-2r)!}$$

Question 3: What is the likelihood that the current year's document uses totally different words from the previous year's?

$$P(r) = {}_{n-r} C_r / {}_n C_r = \frac{(n-r)! (n-r)!}{(n-2r)! n!}$$

Statement: $P(r)$ decreases with r as long as r is less than $(2/3)n$.

Proof: $P(r)/P(r - 1) = (n - 2r + 2)(n - 2r + 1)/(n - r + 1)^2$, where $r \geq 2$. The difference between the numerator and denominator of this ratio is $3r^2 - 2(n + 2)r + (n + 1)$. This difference is a U-shaped function of r with the first intercept with the x -axis less than 1 and the second intercept larger than $(2/3)n$. In other words, as long as $r < (2/3)n$, the difference is negative and therefore the ratio is less than 1 (i.e., $P(\cdot)$ is decreasing in r).

Our setting: (1) we allow a word to be used more than once in a document and (2) we downweight common words used by sample firms as a whole. The decreasing relation between the raw difference score and document length demonstrated above is expected to hold.

TABLE B 1
Estimating the Relation Between the Raw Difference Score and Document Length

| | Coefficient | t-Statistic |
|----------------------------|-------------|-------------|
| <i>Intercept</i> | 0.267 | 57.95 |
| <i>Length</i> | -2.79E-5 | -13.71 |
| <i>Length</i> ² | 1.72E-9 | 6.14 |
| <i>Length</i> ³ | -4.98E-14 | -3.28 |
| <i>Length</i> ⁴ | 6.53E-19 | 1.96 |
| <i>Length</i> ⁵ | -3.13E-24 | -1.30 |
| Model Fit <i>F</i> -stat. | | 417.15 |
| <i>R</i> ² | | 6.9% |

TABLE B 2
Distributions of Rawscore and Score

| Variable | Mean | Std. | Min. | P10 | P25 | Median | P75 | P90 | Max. |
|-----------------|-------|-------|--------|--------|--------|--------|-------|-------|-------|
| <i>Rawscore</i> | 0.155 | 0.147 | 0.000 | 0.028 | 0.053 | 0.106 | 0.207 | 0.354 | 0.978 |
| <i>Score</i> | 0.000 | 0.142 | -0.253 | -0.128 | -0.090 | -0.041 | 0.047 | 0.185 | 0.840 |

2. The procedure to adjust the raw score for document length:

We use a Taylor expansion at 0 to empirically estimate the functional form between the raw difference score and document length. *Rawscore* measures the extent to which the current year's MD&A document differs from the previous year's. *Length* is the number of words in the current year's MD&A. We regress *Rawscore* on the first five polynomials of *Length*. Table B1 reports the OLS estimation.

The fitted scores range between 0.044 and 0.266. The MD&A modification score we use for empirical analyses is the raw score minus the expected score given the document length: $Score = Rawscore - fitted\ score$. Table B2 shows the distributions of *Rawscore* and *Score*.

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